

THE VOTER MODEL IN A RANDOM ENVIRONMENT IN \mathbb{Z}^d

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ABSTRACT. We consider the voter model with flip rates determined by $(\mu_e, e \in E_d)$, where E_d is the set of all non-oriented nearest-neighbour edges in the Euclidean lattice \mathbb{Z}^d . Suppose that $(\mu_e, e \in E_d)$ are i.i.d. random variables satisfying $\mu_e \geq 1$. We prove that when $d = 2$, almost surely for all random environments the voter model has only two extremal invariant measures: δ_0 and δ_1 .

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1. INTRODUCTION

The voter model is an interacting particle system, describing the collective behavior of voters who constantly update their political positions. In this paper, voters are represented by vertices of the Euclidean lattice \mathbb{Z}^d . The voter at x may hold either of two political positions, denoted by 0 or 1. Let $\eta(x)$ be the political position of voter x and the collection $\eta = \{\eta(x); x \in \mathbb{Z}^d\}$ be an element of $\{0, 1\}^{\mathbb{Z}^d}$. The voter at x updates his political position at a random time, following the exponential distribution with parameter $\sum_z \mu_{xz}$, where the summation is over $2d$ nearest neighbors. At the time of update the voter takes the position of his neighbor y with probability $\mu_{xy}/(\sum_z \mu_{xz})$. When $\mu_e \equiv 1$, this is a model well studied in Chapter V of [10].

The voter model can be constructed by the graphical representation, see §3.6 of [10]. This approach not only works for all positive μ_{xy} , but also clearly exhibits the duality relation which will be used in our proof in §3.

We are interested in the case when $(\mu_e, e \in E_d)$ are i.i.d. random variables satisfying $\mu_e \geq 1$, defined on a probability space $(\Omega, \mathcal{F}, \mathbb{P})$. There have been very few literatures about the voter model in random environments. As far as we know, only the one dimensional voter model in a random environment has been explored in [8], and a voter model in the supercritical cluster of the Bernoulli bond percolation is discussed in [3]. As usual, one would like to identify all invariant measures. When all voters take the same position, there

will be no change thereafter. Therefore the configurations that $\eta(x) \equiv 0$ or 1 are traps of the voter model and the measures δ_0 and δ_1 of point mass are invariant. With an extra effort, we are able to identify all invariant measures.

Theorem 1.1. *Let $d = 1$ or 2 . Suppose that (μ_e) are i.i.d. and $\mu_e \geq 1$ \mathbb{P} -a.s. There exists $\Omega_0 \subseteq \Omega$ with $\mathbb{P}(\Omega_0) = 1$. For any $\omega \in \Omega_0$, the voter model has only two extremal invariant measures: δ_0 and δ_1 .*

This extends Corollary 1.13 of Chapter V of [10]. In light of Example 1.5 of Chapter V of [10], the conclusion of the theorem generally does not hold when $d \geq 3$.

The proof of the theorem involves the duality and the dual of the voter model is a coalescing Markov chain taking values on the set of all finite set of vertices of \mathbb{Z}^d . When the initial state is a singleton, the coalescing Markov chain always takes value on singletons. If we identify singleton $\{x\}$ with vertex x , then the coalescing Markov chain is exactly a continuous-time random walk in a random environment. Intuitively, a walker stays at x for an exponential time with parameter μ_x , jumps to a nearest neighbor, say y , with probability $\mu_{xy}/(\sum_z \mu_{xz})$. This is also called the *variable speed random walk* or the *random conductance model*. There is a large amount of literatures on random walks in random environments, e.g., [1], [5] and [6].

Let $\{X_t\}$ and $\{Y_t\}$ be two independent variable speed random walks. The problem we are interested is reduced to the property that $X_t = Y_t$ infinitely often. More specifically, we say $X_t = Y_t$ infinitely often if there exists an infinite sequence of random times $\{t_1, t_2, \dots\}$ with $t_1 < t_2 < \dots$ and $\lim_{i \rightarrow \infty} t_i = \infty$, such that $X(t_i) = Y(t_i)$ for all $i \geq 1$.

Theorem 1.2. *Let $d = 1$ or 2 . Suppose that $(\mu_e, e \in E_d)$ are i.i.d. and $\mu_e \geq 1$ \mathbb{P} -a.s. There exists $\Omega_0 \subseteq \Omega$ with $\mathbb{P}(\Omega_0) = 1$. Let $\omega \in \Omega_0$ and P_ω denote the probability conditional on the environment. If $\{X_t\}$ and $\{Y_t\}$ are two independent variable speed random walks starting from x and y respectively, then $P_\omega(X_t = Y_t \text{ infinitely often}) = 1$.*

The proof for the 2-dimensional case follows a similar idea which is used in [3] and the key to the proof is the heat kernel estimates obtained by Barlow and Deuschel [1]. For more references related to collisions of two random walks, we refer readers to [2] [4] and [9]. Again the conclusion of the theorem holds trivially in dimension one, and fails in general when the dimension is 3 or greater.

This paper is organized as follows. In the next section we quote from [1] the heat kernel bounds as Theorem 2.1, and prove two lemmas that will be used in the proof of the main theorem. Theorems 1.1 and 1.2 will then be

proved in §3. Between the two proofs we will review the coalescing Markov chain and its dual relation with the voter model.

For the typographical reason we write $\mu(x)$ instead of μ_x whenever it is necessary. $[s]$ stands for the largest integer which is smaller than or equal to s . $|A|$ means the cardinality of set A . $\mathbb{Z}^d = (V_d, E_d)$ is the d dimensional Euclidean lattice. Let $|x - y|$ denote the Euclidean distance between x and y . $B_x(s) = \{z : |z - x| \leq s\}$ is the ball of vertices, centered at x . Note that this also makes sense when the radius s is a positive number. Therefore we prefer $B(s)$ over $B([s])$.

2. LEMMAS

Suppose that (μ_e) are i.i.d. and $\mu_e \geq 1$ \mathbb{P} -a.s. Let $\{X_t\}$ be the *variable speed random walk* and $q_t^\omega(x, y)$ be the heat kernel of $\{X_t\}$. Namely,

$$q_t^\omega(x, y) = P_\omega^x(X_t = y).$$

Theorem 2.1. (*Theorem 1.2 of [1]*) *Let $d \geq 2$ and $\sigma \in (0, 1)$. There exist random variables $S_x, x \in \mathbb{Z}^d$, such that*

$$\mathbb{P}(S_x(\omega) \geq n) \leq c_1 \exp(-c_2 n^\sigma), \quad (2.1)$$

and constants c_i (depending only on d and the distribution of μ_e) such that the following hold.

If $|x - y|^2 \vee t \geq S_x^2$, then

$$q_t^\omega(x, y) \leq c_3 t^{-d/2} e^{-c_4 |x-y|^2/t} \text{ when } t \geq |x - y|, \quad (2.2)$$

$$q_t^\omega(x, y) \leq c_3 \exp(-c_4 |x - y|(1 \vee \log(|x - y|/t))) \text{ when } t \leq |x - y|. \quad (2.3)$$

If $t \geq S_x^2 \vee |x - y|^{1+\sigma}$, then

$$q_t^\omega(x, y) \geq c_5 t^{-d/2} e^{-c_6 |x-y|^2/t}. \quad (2.4)$$

Lemma 2.2. *Let $A_n(\omega)$ be the random set defined by*

$$A_n(\omega) = \{x : |x| \leq n, S_x(\omega) \leq 2 \log n\}.$$

Then almost surely there exists a finite random variable $U(\omega)$ such that $|A_n(\omega)| \geq c_7 n^2$ for any $n \geq U(\omega)$.

Proof. Set $W_n(\omega) := |A_n(\omega)| = \sum_{|x| \leq n} \mathbf{1}_{\{S_x(\omega) \leq 2 \log n\}}$. By (2.1),

$$\mathbb{E}W_{2^n} = \sum_{|x| \leq 2^n} (1 - \mathbb{P}(S_x > n \log 4)) \geq \pi 2^{2n} (1 - c_1 \exp\{-c_2 (n \log 4)^\sigma\}).$$

Since $0 \leq W_{2^n} \leq \pi 2^{2n}$, it follows that

$$\mathbb{P}(W_{2^n} < \pi 2^{2n} (1 - c_1 \exp\{-\frac{c_2}{2} (n \log 4)^\sigma\})) \leq \exp\{-\frac{c_2}{2} (n \log 4)^\sigma\}.$$

By the Borel-Cantelli Lemma, $\{W_{2^n} < \pi 2^{2n}(1 - c_1 \exp\{-c_2(n \log 4)^\sigma/2\})\}$ happens finitely many times almost surely. Then there exists $\Omega_0 \subseteq \Omega$ with $\mathbb{P}(\Omega_0) = 1$ and a random variable $U(\omega)$ such that $U(\omega) < \infty$ for all $\omega \in \Omega_0$, and for all $n \geq U(\omega)$,

$$W_{2^n}(\omega) \geq \pi 2^{2n}(1 - c_1 \exp\{-c_2(n \log 4)^\sigma/2\}) \geq c_8 2^{2n} .$$

For all $2^n < k < 2^{n+1}$, $W_k \geq W_{2^n} \geq c_8 2^{2n} \geq c_8 k^2/4$. \square

Lemma 2.3. *Let $\omega \in \Omega_0$, and $\{X_t\}$, $\{Y_t\}$ be independent variable speed random walks in \mathbb{Z}^2 starting from x and y respectively. Then*

$$P_\omega(X_t = Y_t \text{ for some } t \geq 1) \geq \delta > 0,$$

where δ is a constant independent of ω , x and y .

Proof. Fix $\omega \in \Omega_0$. Recall that $B_x(n) = \{z : |z - x| \leq n\}$, and set

$$M_\omega(n) = B_x(n) \cap B_y(n) \cap A_n(\omega) .$$

For $n \geq U(\omega) + (|x| \vee |y|)(1 + 12\pi c_7^{-1})$,

$$\begin{aligned} |M_\omega(n)| &= |B_x(n)| + |B_y(n)| + |A_n(\omega)| - |B_x(n) \cup B_y(n)| - |B_x(n) \cup A_n(\omega)| \\ &\quad - |B_y(n) \cup A_n(\omega)| + |B_x(n) \cup B_y(n) \cup A_n(\omega)| \\ &\geq |B_x(n)| + |B_y(n)| + |A_n(\omega)| - 2|B_0(|x| \vee |y| + n)| \\ &\geq \pi n^2 + \pi n^2 + c_7 n^2 - 2\pi(|x| \vee |y| + n)^2 \\ &= c_7 n^2 - 4\pi n(|x| \vee |y|) - 2\pi(|x| \vee |y|)^2 \\ &\geq c_7 n^2 - 6\pi n(|x| \vee |y|) \geq \frac{c_7}{2} n^2 . \end{aligned} \tag{2.5}$$

Set $T = \exp(\frac{2}{1+\sigma} \log t_0)$, where σ is given in Theorem 2.1, and

$$t_0 = [S_x(\omega) \vee S_y(\omega)]^2 + [U(\omega) + (|x| \vee |y|)(1 + 12\pi c_7^{-1})]^2.$$

Define the random variable

$$H := \int_{t_0}^T \mathbf{1}_{\{X_s = Y_s \in M(s^{1/2})\}} ds .$$

We shall establish a lower bound of $\mathbb{E}H$ and an upper bound of $\mathbb{E}H^2$ below as lemmas.

Lemma 2.4. $\mathbb{E}_\omega H \geq c_9 \log T$.

Lemma 2.5. $\mathbb{E}_\omega H^2 \leq (4\pi c_3^2 + 2\pi^2 c_3^4/c_4)(\log T)^2$.

Then by the Hölder's inequality,

$$P_\omega(H > 0) \geq \frac{(\mathbb{E}_\omega H)^2}{\mathbb{E}_\omega H^2} \geq \frac{(c_9 \log T)^2}{(4\pi c_3^2 + 2\pi^2 c_3^4/c_4)(\log T)^2} = \frac{c_9^2 c_4}{4\pi c_3^2 c_4 + 2\pi^2 c_3^4}.$$

It completes the proof by choosing δ to be $c_9^2 c_4 / (4\pi c_3^2 c_4 + 2\pi^2 c_3^4)$ which is positive. \square

Proof of Lemma 2.4.

$$\begin{aligned} \mathbb{E}_\omega H &= \int_{t_0}^T \mathbb{E}_\omega \mathbf{1}_{\{X_s=Y_s \in M(s^{1/2})\}} ds \\ &= \int_{t_0}^T \sum_{z \in M(s^{1/2})} P_\omega(X_s = z, Y_s = z) ds \\ &= \int_{t_0}^T \sum_{z \in M(s^{1/2})} q_s^\omega(x, z) q_s^\omega(y, z) ds . \end{aligned}$$

Since $z \in M(s^{1/2})$, we have $|x - z|^2 \leq s \leq T = \exp(\frac{2}{1+\sigma} \log t_0)$. Thus $s \geq t_0 \geq S_x^2(\omega) \vee |x - z|^{1+\sigma}$. Similarly $s \geq S_y^2(\omega) \vee |y - z|^{1+\sigma}$. Hence the condition of (2.4) is satisfied, and $\mathbb{E}_\omega H$ can be bounded below.

$$\begin{aligned} \mathbb{E}_\omega H &\geq \int_{t_0}^T \sum_{z \in M(s^{1/2})} c_5^2 s^{-2} \exp\left(-c_6 \frac{|x - z|^2}{s} - c_6 \frac{|y - z|^2}{s}\right) ds \\ &\geq c_5^2 e^{-2c_6} \int_{t_0}^T \sum_{z \in M(s^{1/2})} s^{-2} ds \geq \frac{c_5^2 c_7 e^{-2c_6}}{2} \int_{t_0}^T s^{-1} ds \geq c_9 \log T . \end{aligned}$$

Here we get the first inequality by (2.4), the second inequality by the fact that $|x - z|^2 \leq s$ for $z \in M(s^{1/2})$, and the third inequality by (2.5) that $|M(s^{1/2})| \geq c_7 s/2$ since $s^{1/2} \geq t_0^{1/2} \geq U(\omega) + (|x| \vee |y|)(1 + 12\pi c_7^{-1})$. \square

Proof of Lemma 2.5.

$$\begin{aligned} \mathbb{E}_\omega H^2 &= 2\mathbb{E}_\omega \left(\int_{t_0}^T dt \int_t^T \mathbf{1}_{\{X_t=Y_t \in M(t^{1/2})\}} \mathbf{1}_{\{X_s=Y_s \in M(s^{1/2})\}} ds \right) \\ &= 2 \int_{t_0}^T dt \int_t^T \mathbb{E}_\omega \sum_{z \in M(t^{1/2})} \sum_{w \in M(s^{1/2})} \mathbf{1}_{\{X_t=Y_t=z\}} \mathbf{1}_{\{X_s=Y_s=w\}} ds \\ &= 2 \int_{t_0}^T dt \int_t^T \sum_{z \in M(t^{1/2})} P_\omega^{(x,y)}(X_t = Y_t = z) \sum_{w \in M(s^{1/2})} P_\omega^{(z,z)}(X_{s-t} = Y_{s-t} = w) ds \\ &= 2 \int_{t_0}^T dt \int_t^T \sum_{z \in M(t^{1/2})} q_t^\omega(x, z) q_t^\omega(y, z) \sum_{w \in M(s^{1/2})} q_{s-t}^\omega(z, w) q_{s-t}^\omega(z, w) ds \\ &\leq 2 \int_{t_0}^T dt \left[\sum_{z \in M(t^{1/2})} q_t^\omega(x, z) q_t^\omega(y, z) \int_0^T \sum_{w \in M((s+t)^{1/2})} (q_s^\omega(z, w))^2 ds \right] . \end{aligned}$$

Since $z \in M(t^{1/2})$, $|x - z| \leq t^{1/2} \leq t$. Similarly $|y - z| \leq t$. Moreover $t \geq t_0 \geq [S_x \vee S_y]^2$. Thus the condition of (2.2) is satisfied. Hence

$$q_t^\omega(x, z)q_t^\omega(y, z) \leq c_3^2 t^{-2} .$$

On the other hand, since $z \in M(t^{1/2})$, $z \in A_{t^{1/2}}(\omega)$. Thus by the definition of $A_n(\omega)$ we conclude that $S_z \leq \log t$, and

$$\begin{aligned} & \int_0^T \sum_{w \in M((s+t)^{1/2})} (q_s^\omega(z, w))^2 ds \\ & \leq \log t + \int_{\log t}^T \sum_{w \in B_z(s)} (q_s^\omega(z, w))^2 ds + \int_{\log t}^T \sum_{w \notin B_z(s)} (q_s^\omega(z, w))^2 ds . \end{aligned} \quad (2.6)$$

For $w \notin B_z(s)$, we have $S_z(\omega) \leq \log t \leq s \leq |z - w|$, thus the condition of (2.3) is satisfied. Hence

$$\begin{aligned} \int_{\log t}^T \sum_{w \notin B_z(s)} (q_s^\omega(z, w))^2 ds & \leq \int_{\log t}^T \sum_{w \notin B_z(s)} c_3^2 \exp(-2c_4|z - w|) ds \\ & \leq \int_{\log t}^T \sum_{n=[s]}^{\infty} 2\pi n c_3^2 \exp(-2c_4 n) ds \leq c_{10} . \end{aligned}$$

For $w \in B_z(s)$, $s \geq |z - w|$ and $s \geq S_z(\omega)$, the condition of (2.2) is satisfied. Hence

$$\begin{aligned} & \int_{\log t}^T \sum_{w \in B_z(s)} (q_s^\omega(z, w))^2 ds \\ & \leq \int_{\log t}^T \sum_{w \in B_z(s)} c_3^2 s^{-2} \exp(-2c_4|z - w|^2/s) ds \\ & \leq \int_{\log t}^T [c_3^2 s^{-2} + \sum_{n=1}^{[s]} c_3^2 2\pi n s^{-2} \exp(-2c_4 n^2/s)] ds \\ & \leq c_3^2 \left(\frac{1}{\log t} - \frac{1}{T} \right) + 2\pi c_3^2 \sum_{n=1}^{[T]} n \int_n^T s^{-2} \exp(-2c_4 n^2/s) ds \\ & \leq \frac{c_3^2}{\log t} + 2\pi c_3^2 \sum_{n=1}^{[T]} n \int_{T^{-1}}^{n^{-1}} \exp(-2c_4 n^2 u) du \\ & \leq \frac{c_3^2}{\log t} + \pi \frac{c_3^2}{c_4} \sum_{n=1}^{[T]} n^{-1} \leq \frac{c_3^2}{\log t} + \frac{\pi c_3^2}{c_4} \log T . \end{aligned}$$

Putting two estimates together into (2.6),

$$\int_0^T \sum_w (q_s^\omega(z, w))^2 ds \leq \log t + c_{10} + \frac{c_3^2}{\log t} + \frac{\pi c_3^2}{c_4} \log T \leq (2 + \frac{\pi c_3^2}{c_4}) \log T.$$

Therefore,

$$\begin{aligned} \mathbb{E}_\omega H^2 &\leq 2 \int_{t_0}^T \left(\sum_{z \in M(t^{1/2})} c_3^2 t^{-2} (2 + \frac{\pi c_3^2}{c_4}) \log T \right) dt \\ &\leq 2 \int_{t_0}^T \frac{c_3^2 \pi (2 + \pi c_3^2 / c_4) \log T}{t} dt \leq (4\pi c_3^2 + \frac{2\pi^2 c_3^4}{c_4}) (\log T)^2. \end{aligned}$$

□

3. PROOF OF THEOREMS

Proof of Theorem 1.2. When $d = 1$, (X_t, Y_t) can be viewed as a random walk in the 2-dimensional lattice with random conductance, and is therefore recurrent. In particular (X_t, Y_t) will certainly hit the diagonal line. Namely, $P_\omega(X_t = Y_t, \text{ for some } t > 0) = 1$.

For $d = 2$, let $\delta > 0$ be defined in Lemma 2.3. Fix $\omega \in \Omega_0$. By Lemma 2.3, there exists a function $f : V_2 \times V_2 \mapsto [1, \infty)$, such that for all $x, y \in V_2$,

$$P_\omega^{(x,y)}(X_t = Y_t \text{ for some } 1 < t \leq f(x, y)) \geq \frac{\delta}{2}. \quad (3.1)$$

Set $x_0 = x$, $y_0 = y$ and $t_0 = 0$. Define x_i , y_i and t_i inductively for $i \geq 1$ as follows. Suppose that x_i , y_i and t_i are already defined. Let $\{\tilde{X}_t\}$ and $\{\tilde{Y}_t\}$ be two independent continuous-time random walks starting from x_i and y_i . Define

$$x_{i+1} := \tilde{X}(f(x_i, y_i)), \quad y_{i+1} := \tilde{Y}(f(x_i, y_i)), \quad \text{and } t_{i+1} := t_i + f(x_i, y_i).$$

Define \mathcal{E}_i to be the event that $X_t = Y_t$ for some $t \in (t_i + 1, t_{i+1}]$ for $i \geq 0$. By (3.1) and the strong Markov property,

$$P_\omega(\mathcal{E}_i | X_t, Y_t, t \leq t_i) = P_\omega^{(x_i, y_i)}(\tilde{X}_t = \tilde{Y}_t \text{ for some } 1 < t \leq f(x_i, y_i)) \geq \frac{\delta}{2}.$$

By the second Borel-Cantelli lemma (extended version, see Page 237 of [7]), $P_\omega(\mathcal{E}_i \text{ infinitely often}) = 1$. Furthermore,

$$P_\omega(X_t = Y_t \text{ infinitely often}) \geq P_\omega(\mathcal{E}_i \text{ infinitely often}) = 1.$$

Thus we complete the proof of Theorem 1.2. □

As a dual of the voter model, the coalescing Markov chain $\{A_t\}$ takes value on the set of all finite sets of vertices of \mathbb{Z}^d . Suppose the initial state

A is a finite set of vertices. Construct A_t as follows: Imagine that there is a particle at each $x \in A$. Each particle performs a variable speed random walk starting from a point in A , independent of each other until they meet. Once two particles collide, they coalesce into one particle. Then A_t is the set of locations of all particles at time t .

The process $\{A_t\}$ can also be constructed by the same graphical representation as used in the construction of the voter model. For $\eta \in \{0, 1\}^{\mathbb{Z}^2}$ and a finite set $A \subseteq \mathbb{Z}^2$, define

$$H(\eta, A) = \mathbf{1}_{\{\eta(z)=1 \text{ for all } z \in A\}}.$$

Then the duality relation holds. That is,

$$\mathbb{E}_\omega^\eta H(\eta_t, A) = \mathbb{E}_\omega^A H(\eta, A_t).$$

For the details of the derivation of this equality, see page 230 of [10]. Equivalently,

$$\mathbb{P}_\omega^\eta(\eta_t(x) = 1 \text{ for all } x \in A) = \mathbb{P}_\omega^A(\eta(x) = 1 \text{ for all } x \in A_t). \quad (3.2)$$

Now we are ready to prove that for $\omega \in \Omega_0$, the corresponding voter model has only two extremal invariant measures: δ_0 and δ_1 .

Proof of Theorem 1.1. The following argument is free of dimension. So we will only deal with the 2 dimensional case. Suppose that π is a probability measure on $\{0, 1\}^{\mathbb{Z}^2}$ and π is invariant for the voter model. Let $\alpha(x) = \pi(\eta(x) = 1)$. Then α is a bounded harmonic function. Namely α satisfies

$$\mu_x \alpha(x) = \sum_z \mu_{xz} \alpha(z).$$

The fact that the *variable speed random walk* in \mathbb{Z}^2 is recurrent implies that all bounded harmonic functions are constants. Let $\alpha = \pi(\eta(0) = 1)$. We want to prove by (3.2) that for any finite subset A ,

$$\pi(\eta(x) = 1 \text{ for all } x \in A) = \alpha. \quad (3.3)$$

Let τ be the first time that the coalescing random walk starting from A coalesces into a singleton. By Theorem 1.2, $P_\omega(\tau < \infty) = 1$. If $t > \tau$, then $A_t = \{X_t\}$, where X_t is the variable speed random walk starting from some

point $x \in A$. By (3.2), for any t

$$\begin{aligned}
& \pi(\eta(y) = 1 \text{ for all } y \in A) = \mathbb{P}_\omega^\pi(\eta_t(y) = 1 \text{ for all } y \in A) \\
&= \int \mathbb{P}_\omega^\eta(\eta_t(y) = 1 \text{ for all } y \in A) \pi(d\eta) \\
&= \int \mathbb{P}_\omega^A(\eta(y) = 1 \text{ for all } y \in A_t) \pi(d\eta) \\
&= \int [\mathbb{P}_\omega^A(\eta(y) = 1 \text{ for all } y \in A_t, \tau \leq t) \\
&\quad + \mathbb{P}_\omega^A(\eta(y) = 1 \text{ for all } y \in A_t, \tau > t)] \pi(d\eta) \\
&= \int [\mathbb{P}_\omega^A(\eta(X_t) = 1, \tau \leq t) + \mathbb{P}_\omega^A(\eta(y) = 1 \text{ for all } y \in A_t, \tau > t)] \pi(d\eta) \\
&= \int [\alpha \mathbb{P}_\omega^A(\tau \leq t) + \mathbb{P}_\omega^A(\eta(y) = 1 \text{ for all } y \in A_t, \tau > t)] \pi(d\eta) \\
&= \alpha + \int [-\alpha \mathbb{P}_\omega^A(\tau > t) + \mathbb{P}_\omega^A(\eta(y) = 1 \text{ for all } y \in A_t, \tau > t)] \pi(d\eta).
\end{aligned}$$

Thus

$$|\pi(\eta(y) = 1 \text{ for all } y \in A) - \alpha| \leq 2\mathbb{P}_\omega^A(\tau > t).$$

Letting $t \rightarrow \infty$, we obtain the desired conclusion (3.3), from which we conclude that $\pi = \alpha\delta_1 + (1 - \alpha)\delta_0$. \square

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