## Moments of random sums and Robbins' problem of optimal stopping

Alexander Gnedin<sup>\*</sup> and Alexander Iksanov<sup>†</sup>

## Abstract

Robbins' problem of optimal stopping asks one to minimise the expected *rank* of observation chosen by some nonanticipating stopping rule. We settle a conjecture regarding the *value* of the stopped variable under the rule optimal in the sense of the rank, by embedding the problem in a much more general context of selection problems with the nonanticipation constraint lifted, and with the payoff growing like a power function of the rank.

1. Let  $X_1, \ldots, X_n$  be independent random variables sampled sequentially from the uniform [0, 1] distribution, and let  $Y_1 < \ldots < Y_n$  be their order statistics. The rank  $R_j$  of the variable  $X_j$  is defined by setting  $R_j = k$  on the event  $X_j = Y_k$ . Robbins' problem of optimal stopping [3] asks one to minimize the expected rank  $\mathbb{E}R_{\tau}$  over all stopping times  $\tau$  that assume values in  $\{1, \ldots, n\}$  and are adapted to the natural filtration of the sequence  $X_1, \ldots, X_n$ . Let  $\tau_n$  be the optimal stopping time. The minimum expected rank  $\mathbb{E}R_{\tau_n}$  increases as n grows, and converges to some finite limit v whose exact value is unknown. The closest known upper bound is slightly less than 7/3. Finding v or even improving the existing rough bounds remains a challenge. A major source of difficulties is that the optimal stopping time  $\tau_n$  is a very complicated function of the sample. It seems that  $\tau_n$  has not been computed for n > 3. Moreover, for large n there is no simplification, and the complexity of the optimal stopping time persists in the ' $n = \infty$ ' limiting form of the problem [6].

In a recent paper Bruss and Swan [4] stressed that it is not even known if  $\limsup_n n \mathbb{E} X_{\tau_n}$  is finite. They mentioned that the property was first conjectured in [2]. While the conjecture stems from the attempts to bound v by the comparison with much simpler problem of minimising  $\mathbb{E} X_{\tau}$  (or minor variations of the problem), it seems that the question is of independent interest as a relation between the stopped sample value and its rank. In this note we settle the conjecture by proving a considerably more general assertion:

<sup>\*</sup>Postal address: Department of Mathematics, Utrecht University, Postbus 80010, 3508 TA Utrecht, The Netherlands. E-mail address: A.V.Gnedin@uu.nl

<sup>&</sup>lt;sup>†</sup>Postal address: Faculty of Cybernetics, National T. Shevchenko University of Kiev, Kiev-01033, Ukraine. E-mail address: iksan72@mail.ru

**Proposition 1.** Fix p > 0. For n = 1, 2, ... let  $\sigma_n$  be a random variable with range  $\{1, ..., n\}$  and arbitrary joint distribution with  $X_1, ..., X_n$ . Then

$$\limsup_{n} \mathbb{E}[R_{\sigma_n}]^p < \infty \quad \text{implies} \quad \limsup_{n} n^p \mathbb{E}[X_{\sigma_n}]^p < \infty.$$
(1)

In particular,  $\lim_{n\to\infty} n^p \mathbb{E}[X_{\tau_n}]^p < \infty$  for  $\tau_n$  the stopping time minimising  $\mathbb{E}[R_{\tau}]^p$  over all stopping times adapted to  $X_1, \ldots, X_n$ .

The idea is to bound  $X_{\sigma_n}$  by exploiting properties of a random walk with negative drift.

**2.** Let  $\xi, \xi_1, \xi_2, \ldots$  be iid nonnegative random variables with  $\mu = \mathbb{E}\xi \in (0, \infty)$ . Let  $S_k := \xi_1 + \cdots + \xi_k$  and for  $\lambda > \mu$  let  $M_{\lambda} =: \sup_{k>0} (S_k - \lambda k)$ .

**Proposition 2.** For p > 0

$$\mathbb{E}\xi^{p+1} < \infty \quad \Longleftrightarrow \quad \mathbb{E}M^p_\lambda < \infty.$$

*Proof.* The moment condition on  $\xi$  is equivalent to  $\mathbb{E}[(\xi - \lambda)^+]^{p+1} < \infty$ , and the result follows from Lemma 3.5 in [1].

**Corollary 3.** Suppose  $\mathbb{E}\xi^{p+1} < \infty$  and let  $\sigma$  be a nonnegative integer random variable with  $\mathbb{E}\sigma^p < \infty$ . Then  $\mathbb{E}S^p_{\sigma} < \infty$ .

*Proof.* This follows from  $S^p_{\sigma} \leq (M_{\lambda} + \lambda \sigma)^p \leq c_p (M^p_{\lambda} + \lambda^p \sigma^p)$ , where  $c_p := 2^{p-1} \vee 1$ .  $\Box$ 

**3.** We can apply Corollary 3 to a Poisson-embedded, limiting form of the stopping problem with continuous time [6]. Let  $\xi_1, \xi_2, \ldots$  be iid rate-one exponential variables,  $S_k$  as above, and let  $T_1, T_2, \ldots$  be iid uniform [0, 1] random times, independent of the  $\xi_j$ 's. The points  $(T_k, S_k)$  are the atoms of a homogeneous planar Poisson process  $\mathcal{P}$  in  $[0, 1] \times [0, \infty)$ . To introduce the dynamics, consider an observer whose information at time  $t \in [0, 1]$  is the (infinite) configuration of points of  $\mathcal{P}$  within the strip  $[0, t] \times [0, \infty)$ , that is  $\{(T_k, S_k) : T_k \leq t\}$ . The rank of point  $(T_k, S_k)$  is defined as  $R_{T_k} = k$ , meaning that  $S_k$  is the kth smallest value among  $S_1, S_2, \ldots$ . The piece of information added at time  $T_k$  is the point  $(T_k, S_k)$ , but not the rank  $R_{T_k}$ .

Suppose the objective of the observer is to minimize  $\mathbb{E}[R_{\tau}]^p$  over stopping times  $\tau$  that assume values in the random set  $\{T_1, T_2 \dots\}$  and are adapted to the information flow of the observer. For the optimal stopping time  $\tau_{\infty}$  it is known from the previous studies that  $\mathbb{E}[R_{\tau_{\infty}}]^p < \infty$  (see [6] and [5]). Taking  $\sigma = R_{\tau_{\infty}}$ , we have  $\mathbb{E}[S_{\sigma}]^p < \infty$ . The case p = 1 corresponds to the infinite version of Robbins's problem of minimising the expected rank.

4. To apply the above to a finite sample, we shall use the familiar realisation of uniform order statistics through sums of exponential variables, as

$$(Y_k, 1 \le k \le n) \stackrel{a}{=} (S_k/S_n, 1 \le k \le n).$$

Introducing the event  $A_n := \{n/S_n > 1 + \epsilon\}$ , we can estimate for  $1 \le k \le n$ 

$$n^{p}Y_{k}^{p} = n^{p}Y_{k}^{p}1_{A_{n}} + n^{p}Y_{k}^{p}1_{A_{n}^{c}} \le n^{p}1_{A_{n}} + (1+\epsilon)^{\rho}S_{k}^{p} \le n^{p}1_{A_{n}} + c_{p}(1+\epsilon)^{p}(M_{\lambda}^{p}+\lambda^{p}k^{p}),$$

where we used  $S_k \leq M_{\lambda} + \lambda k$ . Using a large deviation bound for the probability of  $A_n$  and sending  $\epsilon \to 0$  we conclude that for any random variable  $\sigma_n$  with values in  $\{1, \ldots, n\}$ 

$$\limsup_{n} n^{p} \mathbb{E}[Y_{\sigma_{n}}]^{p} \leq c_{p} \lambda^{p} \limsup_{n} \mathbb{E}\sigma_{n}^{p} + c_{p} \mathbb{E}M_{\lambda}^{p}.$$

Finally, taking  $\sigma_n = R_{\tau_n}$ , Proposition 2 follows from

$$\limsup_{n} n^{p} \mathbb{E}[X_{\tau_{n}}]^{p} \leq c_{p} \lambda^{p} \limsup_{n} \mathbb{E}[R_{\tau_{n}}]^{p} + c_{p} \mathbb{E}M_{\lambda}^{p} < \infty,$$

since  $\mathbb{E}[R_{\tau_n}]^p$  converges to a finite limit (see [6], [5]).

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## References

- ALSMEYER, G. AND IKSANOV, A. (2009). A log-type moment result for perpetuities and its application to martingales in supercritical branching random walks. *Elect. J. Probab.* 14, 289–313.
- [2] ASSAF, D. AND SAMUEL-CAHN, E. (1996). The secretary problem: minimizing the expected rank with i.i.d. random variables. *Adv. Appl. Prob.* 28, 828–852.
- [3] BRUSS, F. T. (2005). What is known about Robbins problem? J. Appl. Prob. 42, 108–120.
- [4] BRUSS, F.T. AND SWAN, Y. (2009). A continuous time approach to Robbins' problem of minimizing the expected rank. J. Appl. Prob. 46, 1-18.
- [5] GIANINI, J. AND SAMUELS, S.M. (1976). The infinite secretary problem. Ann. Probab. 3, 418–432.
- [6] GNEDIN, A. (2007). Optimal stopping with rank-dependent loss, J. Appl. Prob. 44, 996–1011.