# Moments of random sums and Robbins' problem of optimal stopping 

Alexander Gnedin* and Alexander Iksanov ${ }^{\dagger}$


#### Abstract

Robbins' problem of optimal stopping asks one to minimise the expected rank of observation chosen by some nonanticipating stopping rule. We settle a conjecture regarding the value of the stopped variable under the rule optimal in the sense of the rank, by embedding the problem in a much more general context of selection problems with the nonanticipation constraint lifted, and with the payoff growing like a power function of the rank.


1. Let $X_{1}, \ldots, X_{n}$ be independent random variables sampled sequentially from the uniform $[0,1]$ distribution, and let $Y_{1}<\ldots<Y_{n}$ be their order statistics. The rank $R_{j}$ of the variable $X_{j}$ is defined by setting $R_{j}=k$ on the event $X_{j}=Y_{k}$. Robbins' problem of optimal stopping [3] asks one to minimize the expected rank $\mathbb{E} R_{\tau}$ over all stopping times $\tau$ that assume values in $\{1, \ldots, n\}$ and are adapted to the natural filtration of the sequence $X_{1}, \ldots, X_{n}$. Let $\tau_{n}$ be the optimal stopping time. The minimum expected rank $\mathbb{E} R_{\tau_{n}}$ increases as $n$ grows, and converges to some finite limit $v$ whose exact value is unknown. The closest known upper bound is slightly less than $7 / 3$. Finding $v$ or even improving the existing rough bounds remains a challenge. A major source of difficulties is that the optimal stopping time $\tau_{n}$ is a very complicated function of the sample. It seems that $\tau_{n}$ has not been computed for $n>3$. Moreover, for large $n$ there is no simplification, and the complexity of the optimal stopping time persists in the ' $n=\infty$ ' limiting form of the problem [6].

In a recent paper Bruss and Swan [4] stressed that it is not even known if $\lim \sup _{n} n \mathbb{E} X_{\tau_{n}}$ is finite. They mentioned that the property was first conjectured in [2]. While the conjecture stems from the attempts to bound $v$ by the comparison with much simpler problem of minimising $\mathbb{E} X_{\tau}$ (or minor variations of the problem), it seems that the question is of independent interest as a relation between the stopped sample value and its rank. In this note we settle the conjecture by proving a considerably more general assertion:

[^0]Proposition 1. Fix $p>0$. For $n=1,2, \ldots$ let $\sigma_{n}$ be a random variable with range $\{1, \ldots, n\}$ and arbitrary joint distribution with $X_{1}, \ldots, X_{n}$. Then

$$
\begin{equation*}
\limsup _{n} \mathbb{E}\left[R_{\sigma_{n}}\right]^{p}<\infty \quad \text { implies } \quad \underset{n}{\limsup } n^{p} \mathbb{E}\left[X_{\sigma_{n}}\right]^{p}<\infty \tag{1}
\end{equation*}
$$

In particular, $\lim _{n \rightarrow \infty} n^{p} \mathbb{E}\left[X_{\tau_{n}}\right]^{p}<\infty$ for $\tau_{n}$ the stopping time minimising $\mathbb{E}\left[R_{\tau}\right]^{p}$ over all stopping times adapted to $X_{1}, \ldots, X_{n}$.

The idea is to bound $X_{\sigma_{n}}$ by exploiting properties of a random walk with negative drift.
2. Let $\xi, \xi_{1}, \xi_{2}, \ldots$ be iid nonnegative random variables with $\mu=\mathbb{E} \xi \in(0, \infty)$. Let $S_{k}:=\xi_{1}+\cdots+\xi_{k}$ and for $\lambda>\mu$ let $M_{\lambda}=: \sup _{k \geq 0}\left(S_{k}-\lambda k\right)$.
Proposition 2. For $p>0$

$$
\mathbb{E} \xi^{p+1}<\infty \quad \Longleftrightarrow \quad \mathbb{E} M_{\lambda}^{p}<\infty
$$

Proof. The moment condition on $\xi$ is equivalent to $\mathbb{E}\left[(\xi-\lambda)^{+}\right]^{p+1}<\infty$, and the result follows from Lemma 3.5 in [1].

Corollary 3. Suppose $\mathbb{E} \xi^{p+1}<\infty$ and let $\sigma$ be a nonnegative integer random variable with $\mathbb{E} \sigma^{p}<\infty$. Then $\mathbb{E} S_{\sigma}^{p}<\infty$.

Proof. This follows from $S_{\sigma}^{p} \leq\left(M_{\lambda}+\lambda \sigma\right)^{p} \leq c_{p}\left(M_{\lambda}^{p}+\lambda^{p} \sigma^{p}\right)$, where $c_{p}:=2^{p-1} \vee 1$.
3. We can apply Corollary 3 to a Poisson-embedded, limiting form of the stopping problem with continuous time [6]. Let $\xi_{1}, \xi_{2}, \ldots$ be iid rate-one exponential variables, $S_{k}$ as above, and let $T_{1}, T_{2}, \ldots$ be iid uniform $[0,1]$ random times, independent of the $\xi_{j}$ 's. The points $\left(T_{k}, S_{k}\right)$ are the atoms of a homogeneous planar Poisson process $\mathcal{P}$ in $[0,1] \times[0, \infty)$. To introduce the dynamics, consider an observer whose information at time $t \in[0,1]$ is the (infinite) configuration of points of $\mathcal{P}$ within the strip $[0, t] \times$ $[0, \infty)$, that is $\left\{\left(T_{k}, S_{k}\right): T_{k} \leq t\right\}$. The rank of point $\left(T_{k}, S_{k}\right)$ is defined as $R_{T_{k}}=k$, meaning that $S_{k}$ is the $k$ th smallest value among $S_{1}, S_{2}, \ldots$. The piece of information added at time $T_{k}$ is the point $\left(T_{k}, S_{k}\right)$, but not the rank $R_{T_{k}}$.

Suppose the objective of the observer is to minimize $\mathbb{E}\left[R_{\tau}\right]^{p}$ over stopping times $\tau$ that assume values in the random set $\left\{T_{1}, T_{2} \ldots\right\}$ and are adapted to the information flow of the observer. For the optimal stopping time $\tau_{\infty}$ it is known from the previous studies that $\mathbb{E}\left[R_{\tau_{\infty}}\right]^{p}<\infty$ (see [6] and [5]). Taking $\sigma=R_{\tau_{\infty}}$, we have $\mathbb{E}\left[S_{\sigma}\right]^{p}<\infty$. The case $p=1$ corresponds to the infinite version of Robbins's problem of minimising the expected rank.
4. To apply the above to a finite sample, we shall use the familiar realisation of uniform order statistics through sums of exponential variables, as

$$
\left(Y_{k}, 1 \leq k \leq n\right) \stackrel{d}{=}\left(S_{k} / S_{n}, 1 \leq k \leq n\right) .
$$

Introducing the event $A_{n}:=\left\{n / S_{n}>1+\epsilon\right\}$, we can estimate for $1 \leq k \leq n$
$n^{p} Y_{k}^{p}=n^{p} Y_{k}^{p} 1_{A_{n}}+n^{p} Y_{k}^{p} 1_{A_{n}^{c}} \leq n^{p} 1_{A_{n}}+(1+\epsilon)^{\rho} S_{k}^{p} \leq n^{p} 1_{A_{n}}+c_{p}(1+\epsilon)^{p}\left(M_{\lambda}^{p}+\lambda^{p} k^{p}\right)$,
where we used $S_{k} \leq M_{\lambda}+\lambda k$. Using a large deviation bound for the probability of $A_{n}$ and sending $\epsilon \rightarrow 0$ we conclude that for any random variable $\sigma_{n}$ with values in $\{1, \ldots, n\}$

$$
\limsup _{n} n^{p} \mathbb{E}\left[Y_{\sigma_{n}}\right]^{p} \leq c_{p} \lambda^{p} \limsup _{n} \mathbb{E} \sigma_{n}^{p}+c_{p} \mathbb{E} M_{\lambda}^{p}
$$

Finally, taking $\sigma_{n}=R_{\tau_{n}}$, Proposition 2 follows from

$$
\limsup _{n} n^{p} \mathbb{E}\left[X_{\tau_{n}}\right]^{p} \leq c_{p} \lambda^{p} \limsup _{n} \mathbb{E}\left[R_{\tau_{n}}\right]^{p}+c_{p} \mathbb{E} M_{\lambda}^{p}<\infty,
$$

since $\mathbb{E}\left[R_{\tau_{n}}\right]^{p}$ converges to a finite limit (see [6], [5]).
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[^0]:    *Postal address: Department of Mathematics, Utrecht University, Postbus 80010, 3508 TA Utrecht, The Netherlands. E-mail address: A.V.Gnedin@uu.nl
    ${ }^{\dagger}$ Postal address: Faculty of Cybernetics, National T. Shevchenko University of Kiev, Kiev-01033, Ukraine. E-mail address: iksan72@mail.ru

