

# $t$ -Wise Independence with Local Dependencies

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## Abstract

In this note we prove a large deviation bound on the sum of random variables with the following dependency structure: there is a dependency graph  $G$  with a bounded chromatic number, in which each vertex represents a random variable. Variables that are represented by neighboring vertices may be arbitrarily dependent, but collections of variables that form an independent set in  $G$  are  $t$ -wise independent.

## 1 Introduction

It is often useful to consider a random variable  $X = \sum_{i=1}^n X_i$  and bound the probability that such a sum deviates from its expectation. For independent  $X_i$ 's, famous bounds are those of Chernoff [3] and Hoeffding [6].

Sums of variables that are not fully independent but have some sort of a dependency structure have also been studied – see for example Pemmaraju [10] or the survey of Janson and Ruciński [8]. Here we are interested in a setting studied by Janson [7]: roughly, in his formulation there is a dependency graph  $G$  of  $n$  vertices, in which each variable is represented by a vertex. Two variables whose corresponding vertices are connected by an edge may be dependent, whereas independent sets of the graph are independent (see Section 2 for a more formal description). Janson provides several applications for his bound, such as U-statistics and the existence of long patterns in random strings (see [7] for more details).

A different form of dependency structure is motivated by the computer science literature. This is the setting in which the  $X_i$ 's are  $t$ -wise independent (see for example Bellare and Rompel [1]). This means that every set of  $t$  variables is independent, but any  $t + 1$  may not be.

In this note we consider the situation in which the random variables  $X_i$  are dependent in both fashions: on the one hand, their dependencies are described by a dependency graph  $G$ . On the other hand, variables represented by independent sets of the graph are not fully independent, but only  $t$ -wise independent. We combine standard techniques used in tail bounds for sums of  $t$ -wise independent random variables with the technique of Janson [7], and in this manner obtain a tail bound for sums of random variables that are both  $t$ -wise independent and have local dependencies.

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## 1.1 Motivation

The tail bounds we prove seem to be applicable in numerous situations, and we will now state two such examples. First, our inequality was used in a recent game theoretic work of Gradwohl and Reingold [5]. Second, the inequality can be used in the hidden pattern problem as a bound on the number of patterns in a random string that is not fully independent.

### 1.1.1 Game theory

In a recent work in game theory, Kalai [9] showed that in certain types of large Bayesian games, the Nash equilibria are not affected by such details as order of play, possibility of revision, and more. One of the necessary assumptions in obtaining this result is the independence of certain random variables related to the players. This was necessary because of the repeated application of a Chernoff bound.

In a generalization of this work, Gradwohl and Reingold [5] showed how to replace this assumption of independence by a more general one of limited correlation. One of the tools used by [5] is Corollary 3.2.

### 1.1.2 Hidden Pattern Problem

In the hidden pattern problem, we are given a sequence of  $n$  random letters from a finite alphabet  $A$ , say  $X_1, \dots, X_n$ . Given a word of fixed length  $d$ , say  $w \in A^d$ , one seeks the number of subsequences  $i_1 < \dots < i_d$  such that  $X_{i_1} \circ \dots \circ X_{i_d} = w$ . The case in which the  $X_i$ 's are independent was studied by Flajolet et al. as well as Janson [7]. Bourdon and Vallée generalize the work by considering strings  $X_1, \dots, X_n$  in which the  $X_i$ 's are not fully independent, but rather are generated by dynamical sources (see [2]). Theorem 3.1 can be used in a straightforward manner to obtain a result similar to that of [7], but applied to variables that are  $t$ -wise independent.

## 2 Definitions

Let  $n \in \mathbb{N}$  be an integer. We denote  $[n] = \{1, \dots, n\}$ . For a graph  $G$ , we denote by  $V(G)$  the vertex set of  $G$ , and by  $E(G)$  the edge set of  $G$  (we will consider only simple undirected graphs). Let  $G$  be a graph of size  $|V(G)| = n$ . We usually think of  $V(G)$  as  $[n]$ . The following three definitions are standard graph definitions.

**Definition 2.1 (independent set)**  $S \subseteq V(G)$  is an independent set of vertices in  $G$  if no two vertices in  $S$  share an edge (according to  $G$ ).

**Definition 2.2 (coloring)** For  $k \in \mathbb{N}$ , a  $k$ -coloring of  $G$  is a map from  $V(G)$  to  $[k]$  such that each two adjacent vertices are mapped to different integers.

**Definition 2.3 (chromatic number)** The chromatic number of  $G$ , denoted by  $\chi(G)$ , is the smallest integer  $k$  such that there exists a  $k$ -coloring of  $G$ .

Note that if the degree of  $G$  is at most  $d \in \mathbb{N}$ , then  $\chi(G) \leq d + 1$  (since the greedy algorithm for coloring works).

Next, we give three definitions concerning distributions: the first one is the standard definition of  $t$ -wise independence, the second definition is of a dependency graph, and the third definition, which combines the first two definitions, is of the family of distributions for which our tail bounds apply.

**Definition 2.4 ( $t$ -wise independence)** For  $m, t \in \mathbb{N}$ , the random variables  $Y_1, \dots, Y_m$  are  $t$ -wise independent, if for every  $T \subseteq [m]$  of size  $t$  the set of variables  $\{Y_i : i \in T\}$  is independent.

**Definition 2.5 (agree)** Let  $n \in \mathbb{N}$ , and let  $G$  be a graph of size  $n$ . We say that the random variables  $X_1, \dots, X_n$  agree with the graph  $G$ , if for every independent set of vertices  $S \subseteq V(G)$ , the set of variables  $\{X_i : i \in S\}$  is independent ( $G$  is sometimes called a dependency graph).

**Definition 2.6 ( $t$ -agree)** For a dependency graph  $G$  as above, we say that the random variables  $X_1, \dots, X_n$   $t$ -agree with  $G$ , if for every independent set of vertices  $S \subseteq V(G)$ , the set of variables  $\{X_i : i \in S\}$  is  $t$ -wise independent.

### 3 Results

We are now ready to state our main result, which is a large deviation bound on the sum of random variables that  $t$ -agree with a graph  $G$  of chromatic number  $\chi(G)$ .

**Theorem 3.1** Let  $n, t \in \mathbb{N}$  be such that  $t > 0$  is even. Let  $G$  be a graph of size  $n$ , and let  $X_1, \dots, X_n$  be random variables that take values in  $[0, 1]$  and  $t$ -agree with  $G$ . Let  $X = \sum_{i \in [n]} X_i$  and let  $\mu = \mathbb{E}[X]$ . Then, for every positive real  $a > 0$ ,

$$\mathbb{P}[|X - \mu| \geq a] < 2\sqrt{\pi t} \cdot \left( \frac{\sqrt{nt \cdot \chi(G)}}{a} \right)^t.$$

When the random variables are Bernoulli and the graph is of bounded degree we have the following corollary:

**Corollary 3.2** Let  $n, d, t \in \mathbb{N}$  be such that  $t > 0$  is even. Let  $G$  be a graph of size  $n$  and degree at most  $d$ . Let  $p \in (0, 1)$ , and let  $X_1, \dots, X_n$  be  $Be(p)$  random variables that  $t$ -agree with  $G$ . Let  $X = \sum_{i \in [n]} X_i$ . Then for every positive real  $a > 0$ ,

$$\mathbb{P}[X \geq (1 + a)pn] < 2\sqrt{\pi t} \cdot \left( \frac{\sqrt{(d+1) \cdot t}}{ap\sqrt{n}} \right)^t.$$

The same bound holds for  $\mathbb{P}[X \leq (1 - a)pn]$ .

### 4 Proof of Main Result

In our proof we will need to bound the  $t$ -moment of the sum of  $t$ -wise independent random variables. The following bound is well known – see Bellare and Rompel [1] for a proof.

**Lemma 4.1** *Let  $m, t \in \mathbb{N}$  be such that  $t > 0$  is even, and let  $Y_1, \dots, Y_m$  be  $t$ -wise independent random variables taking values in  $[0, 1]$ . Let  $Y = \sum_{i \in [m]} Y_i$  and let  $\mu = \mathbb{E}[Y]$ . Then*

$$\mathbb{E}[(Y - \mu)^t] < 2 \cdot e^{\frac{1}{6t}} \cdot \sqrt{\pi t} \cdot \left(\frac{mt}{e}\right)^{t/2}.$$

We now prove Theorem 3.1.

**Proof:** Let  $G$  be a graph of size  $n$ , and let  $X_1, \dots, X_n$  be random variables that  $t$ -agree with  $G$ . Let  $X = \sum_{i \in [n]} X_i$  and let  $\mu = \mathbb{E}[X]$ . Let  $f$  be a  $k$ -coloring of  $G$  such that  $k = \chi(G)$ . For every  $j \in [k]$ , denote

$$V_j = f^{-1}(j),$$

which is an independent set of vertices. So for all  $j \in [k]$ , the set of variables  $\{X_i : i \in V_j\}$  is  $t$ -wise independent. By Lemma 4.1, for every  $j \in [k]$ ,

$$\mathbb{E}[(Y_j - \mu_j)^t] < 2 \cdot e^{\frac{1}{6t}} \cdot \sqrt{\pi t} \cdot \left(\frac{|V_j|t}{e}\right)^{t/2},$$

where  $Y_j = \sum_{i \in V_j} X_i$  and  $\mu_j = \mathbb{E}[Y_j]$ .

We now bound the  $t$ -moment of  $X$ . Let  $p_1, \dots, p_k$  be  $k$  non-negative real numbers such that  $\sum_{j \in [k]} p_j = 1$  (to be determined later). By Jensen's inequality and linearity of expectation,

$$\begin{aligned} \mathbb{E}[(X - \mu)^t] &= \mathbb{E}\left[\left(\sum_{j \in [k]} p_j \frac{Y_j - \mu_j}{p_j}\right)^t\right] \\ &\leq \sum_{j \in [k]} p_j \frac{\mathbb{E}[(Y_j - \mu_j)^t]}{p_j^t} \\ &< \sum_{j \in [k]} p_j \frac{2 \cdot e^{\frac{1}{6t}} \cdot \sqrt{\pi t} \cdot \left(\frac{|V_j|t}{e}\right)^{t/2}}{p_j^t}. \end{aligned}$$

For  $j \in [k]$ , set  $q_j \in \mathbb{R}$  to be such that

$$q_j^t = 2 \cdot e^{\frac{1}{6t}} \cdot \sqrt{\pi t} \cdot \left(\frac{|V_j|t}{e}\right)^{t/2},$$

and set

$$p_j = \frac{q_j}{\sum_{\ell \in [k]} q_\ell}.$$

Substituting these values of  $p_j$  yields

$$\mathbb{E}[(X - \mu)^t] < \left(\sum_{j \in [k]} q_j\right)^t = \left(\sum_{j \in [k]} (2 \cdot e^{\frac{1}{6t}} \cdot \sqrt{\pi t})^{1/t} \cdot \left(\frac{|V_j|t}{e}\right)^{1/2}\right)^t,$$

which implies that

$$\mathbb{E} [(X - \mu)^t] < 2 \cdot \sqrt{\pi t} \cdot (\sqrt{knt})^t$$

by Cauchy-Schwartz. Since  $t$  is even we now use Markov's Inequality, implying that for every real number  $a > 0$ ,

$$\mathbb{P} [|X - \mu| \geq a] = \mathbb{P} [(X - \mu)^t \geq a^t] \leq \frac{\mathbb{E} [(X - \mu)^t]}{a^t} < 2\sqrt{\pi t} \cdot \left(\frac{\sqrt{knt}}{a}\right)^t.$$

Substituting  $k = \chi(G)$ , we get that

$$\mathbb{P} [|X - \mu| \geq a] < 2\sqrt{\pi t} \cdot \left(\frac{\sqrt{nt \cdot \chi(G)}}{a}\right)^t.$$

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