

# Interpolating Thin-Shell and Sharp Large-Deviation Estimates For Isotropic Log-Concave Measures

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## Abstract

Given an isotropic random vector  $X$  with log-concave density in Euclidean space  $\mathbb{R}^n$ , we study the concentration properties of  $|X|$ . We show in particular that:

$$\mathbb{P}(|X| \geq (1+t)\sqrt{n}) \leq \exp(-cn^{\frac{1}{2}} \min(t^3, t)) \quad \forall t > 0,$$

for some universal constant  $c > 0$ . This improves the best known deviation results above the expectation on the thin-shell and mesoscopic scales due to Fleury and Klartag, respectively, and recovers the sharp large-deviation estimate of Paouris. Another new feature of our estimate is that it improves when  $X$  is  $\psi_\alpha$  ( $\alpha \in (1, 2]$ ), in precise agreement with the sharp Paouris estimates. The upper bound on the thin-shell width  $\sqrt{\text{Var}(|X|)}$  we obtain is of the order of  $n^{1/3}$ , and improves down to  $n^{1/4}$  when  $X$  is  $\psi_2$ . Our estimates thus continuously interpolate between a new best known thin-shell estimate and the sharp Paouris large-deviation one.

## 1 Introduction

Let a Euclidean norm  $|\cdot|$  on  $\mathbb{R}^n$  be fixed. This work is dedicated to quantitative concentration properties of  $|X|$ , where  $X$  is an isotropic random vector in  $\mathbb{R}^n$  with log-concave density. Recall that a random vector  $X$  in  $\mathbb{R}^n$  (and its density) is called isotropic if  $\mathbb{E}X = 0$  and  $\mathbb{E}X \otimes X = Id$ , i.e. its barycenter is at the origin and its covariance matrix is equal to the identity one. Taking traces, we observe that  $\mathbb{E}|X|^2 = n$ . Here and throughout we use  $\mathbb{E}$  to denote expectation,  $\mathbb{P}$  to denote probability, and  $\text{Var}$  to denote variance. A function  $g : \mathbb{R}^n \rightarrow \mathbb{R}_+$  is called log-concave if  $-\log g : \mathbb{R}^n \rightarrow \mathbb{R} \cup \{+\infty\}$  is convex. Throughout this work,  $C, c, c_2, C'$ , etc. denote universal positive numeric constants, independent of any other parameter and in particular the dimension  $n$ , whose value may change from one occurrence to the next.

It was conjectured by Anttila, Ball and Perissinaki [1] that  $|X|$  is concentrated around its expectation significantly more than suggested by the trivial bound  $\text{Var}|X| \leq \mathbb{E}|X|^2 =$

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$n$ . Namely, they conjectured that there exists a sequence  $\{\varepsilon_n\}$  decreasing to 0 with the dimension  $n$ , so that  $X$  is concentrated within a “thin shell” of relative width  $2\varepsilon_n$  around the (approximately) expected Euclidean norm of  $\sqrt{n}$ :

$$\mathbb{P}(|X| - \sqrt{n} \geq \varepsilon_n \sqrt{n}) \leq \varepsilon_n . \quad (1.1)$$

Their conjecture was mainly motivated by the Central Limit Problem for log-concave measures, and as pointed out in [1], implies that most marginals of log-concave measures are approximately Gaussian.

A stronger version of this conjecture was put forth by Bobkov and Koldobsky [7]. It may be equivalently formulated as stating that the “thin-shell width”  $\sqrt{\text{Var}|X|}$  is bounded above by a universal constant  $C$ .

An even stronger conjecture is due to Kannan, Lovász and Simonovits [16]. In an equivalent form, it states that for any smooth function  $f : \mathbb{R}^n \rightarrow \mathbb{R}$ :

$$\text{Var}(f(X)) \leq C \mathbb{E}|\nabla f(X)|^2 .$$

Applied to the function  $f(x) = |x|^p$  with  $p = c\sqrt{n}$ , the KLS conjecture implies (see [11] and Section 4) that:

$$\mathbb{P}(|X| - \sqrt{n} \geq t\sqrt{n}) \leq C \exp(-c\sqrt{nt}) \quad \forall t \geq 0 . \quad (1.2)$$

It was shown by G. Paouris [29] that the predicted positive deviation estimate (1.2) indeed holds in the large:

$$\mathbb{P}(|X| \geq (1+t)\sqrt{n}) \leq \exp(-c\sqrt{nt}) \quad \forall t \geq C > 0 . \quad (1.3)$$

Moreover, Paouris showed that when  $X$  is  $\psi_\alpha$  ( $\alpha \in [1, 2]$ ):

$$\mathbb{P}(|X| \geq (1+t)\sqrt{n}) \leq \exp(-cn^{\frac{\alpha}{2}}t) \quad \forall t \geq C > 0 . \quad (1.4)$$

Recall that  $X$  (and its density) is said to be “ $\psi_\alpha$  with constant  $D > 0$ ” if:

$$(\mathbb{E}|\langle X, y \rangle|^p)^{1/p} \leq Dp^{1/\alpha} \left( \mathbb{E}|\langle X, y \rangle|^2 \right)^{1/2} \quad \forall p \geq 2 \quad \forall y \in \mathbb{R}^n .$$

We will simply say that “ $X$  is  $\psi_\alpha$ ”, if it is  $\psi_\alpha$  with constant  $D \leq C$ , and not specify explicitly the dependence of the estimates on the parameter  $D$ . By Borell’s Lemma [10] (see also [27, Appendix III]), it is well known that any  $X$  with log-concave density is  $\psi_1$  with some universal constant, and so we only gain additional information when  $\alpha > 1$ .

The positive large-deviation estimate (1.4) is easily verified to be sharp (up to universal constants) for all  $\alpha \in [1, 2]$ . However, this leaves open the concentration estimates in the bulk: positive deviation  $\mathbb{P}(|X| \geq (1+t)\sqrt{n})$  when  $t \in [0, C]$ , and negative deviation  $\mathbb{P}(|X| \leq (1-t)\sqrt{n})$  when  $t \in [0, c]$  ( $c \in (0, 1)$ ); in particular, this gives no information on the thin-shell  $\sqrt{\text{Var}|X|}$ . We remark that the small-ball estimates  $\mathbb{P}(|X| \leq \varepsilon\sqrt{n})$  for  $\varepsilon \in [0, 1-c]$  will mostly be disregarded in this work (see however Remark 1.4 below).

The first non-trivial estimate on the concentration of  $|X|$  around its expectation was given by B. Klartag in [18], involving logarithmic improvements in  $n$  over the trivial bounds. This validated the conjectured thin-shell concentration (1.1), allowing Klartag to resolve the Central Limit Problem for log-concave measures. A different proof continuing Paouris' approach was given by Fleury, Guédon and Paouris in [13]. Klartag then improved in [19] his estimates from logarithmic to polynomial in  $n$  as follows (for any small  $\varepsilon > 0$ ):

$$\mathbb{P}(|X| - \sqrt{n} \geq t\sqrt{n}) \leq C_\varepsilon \exp(-c_\varepsilon n^{\frac{1}{3}-\varepsilon} t^{\frac{10}{3}-\varepsilon}) \quad \forall t \in [0, 1]. \quad (1.5)$$

This implies in particular a thin-shell estimate of:

$$\sqrt{\text{Var}|X|} \leq C_\varepsilon n^{\frac{1}{2}-\frac{1}{10}-\varepsilon}.$$

Note, however, that when  $t = 1/2$ , (1.5) does not recover the sharp positive large-deviation estimate of Paouris (1.3).

Recently in [12], B. Fleury improved Klartag's thin-shell estimate to:

$$\sqrt{\text{Var}|X|} \leq C n^{\frac{1}{2}-\frac{1}{8}},$$

by obtaining the following deviation estimates:

$$\mathbb{P}(|X| \geq (1+t)\sqrt{n}) \leq C \exp(-cn^{\frac{1}{4}}t^2) \quad \forall t \in [0, 1];$$

$$\mathbb{P}(|X| \leq (1-t)\sqrt{n}) \leq C \exp(-cn^{\frac{1}{8}}t) \quad \forall t \in [0, 1].$$

Note, however, that when  $t = 1/2$ , Fleury's positive and negative large-deviation estimates are both inferior to those of Klartag, and so in the mesoscopic scale  $t = n^{-\delta}$  ( $\delta > 0$  small), Klartag's estimates still outperform Fleury's (and Paouris' ones are inapplicable). In addition, note that both Klartag and Fleury's estimates do not seem to improve under a  $\psi_\alpha$  condition, contrary to the ones of Paouris. See also [8, 20, 11, 23] for further related results.

All of this suggests that one might hope for a concentration estimate which:

- Recovers the sharp positive large-deviation result of Paouris (1.4).
- Improves if  $X$  is  $\psi_\alpha$ .
- Improves the best-known thin-shell estimate of Fleury.
- Improves the best-known positive mesoscopic-deviation estimate of Klartag.
- Interpolates continuously between all positive scales of  $t$  (bulk, mesoscopic, large-deviation).

The aim of this work is to provide precisely such an estimate.

## 1.1 The Results

**Theorem 1.1.** *Let  $X$  denote an isotropic random vector in  $\mathbb{R}^n$  with log-concave density, which is in addition  $\psi_\alpha$  ( $\alpha \in [1, 2]$ ). Then:*

$$\mathbb{P}(|X| \geq (1+t)\sqrt{n}) \leq \exp(-cn^{\frac{\alpha}{2}} \min(t^{2+\alpha}, t)) \quad \forall t > 0, \quad (1.6)$$

and:

$$\mathbb{P}(|X| \leq (1-t)\sqrt{n}) \leq C \exp(-cn^{\frac{\alpha}{2(2+\alpha)}} t) \quad \forall t \in (0, 1). \quad (1.7)$$

In particular:

$$\sqrt{\text{Var}(|X|)} \leq Cn^{\frac{1}{2+\alpha}}. \quad (1.8)$$

Note that when  $\alpha = 1$ , as is the case for an arbitrary isotropic  $X$  with log-concave density, we obtain the following thin-shell estimate:

$$\sqrt{\text{Var}(|X|)} \leq Cn^{\frac{1}{2} - \frac{1}{6}}.$$

Also note that we obtain  $\mathbb{P}(|X| \geq (1+\varepsilon)\sqrt{n}) \leq \exp(-C_\varepsilon n^{\frac{\alpha}{2}})$  for any  $\varepsilon > 0$ , whereas Paouris' estimate (1.4) only ensures that this holds for  $\varepsilon \geq C$  for some large enough  $C > 0$ .

Theorem 1.1 is a standard consequence of the following moment estimates, which are the main result of this work:

**Theorem 1.2.** *Let  $X$  denote an isotropic random vector in  $\mathbb{R}^n$  with log-concave density, which is in addition  $\psi_\alpha$  ( $\alpha \in [1, 2]$ ). Then for any  $2 \leq p \leq cn^{\alpha/2}$ :*

$$(\mathbb{E}|X|^p)^{\frac{1}{p}} \leq \left(1 + C \left(\frac{p}{n^{\frac{\alpha}{2}}}\right)^{\frac{1}{\alpha+1}}\right) (\mathbb{E}|X|^2)^{\frac{1}{2}}. \quad (1.9)$$

Note that using  $p = 4$  in (1.9) we obtain  $\text{Var}(|X|^2) \leq C_2 n^{2 - \frac{\alpha}{2(\alpha+1)}}$ , which only yields  $\sqrt{\text{Var}(|X|)} \leq C_3 n^{\frac{1}{2} - \frac{\alpha}{4(\alpha+1)}}$ , an inferior estimate to (1.8). The reason for this discrepancy is due to the fact that our moment estimates for relatively small values of  $p$  may be a-posteriori improved, by integrating by parts and using the deviation estimates of Theorem 1.1 (see [11, Lemma 6]):

**Corollary 1.3.** *With the same assumptions as in Theorem 1.2, for any  $2 \leq p \leq c_2 n^{\frac{\alpha}{2(\alpha+2)}}$ :*

$$(\mathbb{E}|X|^p)^{\frac{1}{p}} \leq \left(1 + C_2 \frac{p}{n^{\frac{\alpha}{\alpha+2}}}\right) (\mathbb{E}|X|^2)^{\frac{1}{2}}.$$

**Remark 1.4.** Note that our estimates for negative mesoscopic and large deviation are still inferior to those of Klartag. The best known negative large-deviation and small-ball estimates are due to Paouris [30], who showed that there exists a constant  $C > 1$  so that:

$$\mathbb{P}(|X| \leq \varepsilon\sqrt{n}) \leq (C\varepsilon)^{cn^{\frac{\alpha}{2}}} \quad \forall \varepsilon \in (0, 1/C). \quad (1.10)$$

It should be possible to extend our methods to handle negative moments  $p$  in Theorem 1.2, resulting in a continuous interpolation between (1.7) and (1.10), and thus improving over Klartag's negative mesoscopic deviation estimates; we leave this for another note.

## 1.2 The Approach

We let  $G_{n,k}$  denote the Grassmann manifold of all  $k$ -dimensional linear subspaces of  $\mathbb{R}^n$ , and  $SO(n)$  the group of rotations. Fixing a Euclidean structure on  $\mathbb{R}^n$ , and given a linear subspace  $F$ , we denote by  $S(F)$  and  $B_2(F)$  the unit-sphere and unit-ball in  $F$ , respectively. When  $F = \mathbb{R}^n$ , we simply write  $S^{n-1}$  and  $B_2^n$ . We denote by  $P_F$  the orthogonal projection onto  $F$  in  $\mathbb{R}^n$ , and given a random vector  $Y$  with density  $g$ , we denote by  $\pi_F g$  the marginal density of  $g$  on  $F$ , i.e. the density of  $P_F Y$ . When  $F = \text{span}(\theta)$ ,  $\theta \in S^{n-1}$ , we denote the density  $\pi_\theta g$  on  $\mathbb{R}$  given by  $\pi_\theta g(t) := \pi_F g(t\theta)$ .

For the proof of Theorem 1.2, we use many of the ingredients developed previously by Klartag [19], and adapted to the language of moments by Fleury [11, 12]:

- It is enough to verify (1.9) with  $X$  replaced by  $Y = (X + G_n)/\sqrt{2}$ , where  $G_n$  denotes a standard Gaussian random vector in  $\mathbb{R}^n$ .
- It is useful to first project  $Y$  onto a lower-dimensional subspace  $F \in G_{n,k}$ . This idea also appears in essence in the work of Paouris [29]. Klartag and Paouris use V. Milman's approach to Dvoretzky's theorem [25, 27] for identifying lower-dimensional structures in most marginals  $P_F Y$ . Fleury, on the other hand, takes an average over the Haar measure on  $G_{n,k}$ , which is more efficient (see [12] or below):

$$\frac{(\mathbb{E}|Y|^p)^{1/p}}{(\mathbb{E}|Y|^2)^{1/2}} \leq \frac{(\mathbb{E}_{F,Y}|P_F Y|^p)^{1/p}}{(\mathbb{E}_{F,Y}|P_F Y|^2)^{1/2}}. \quad (1.11)$$

- Rewriting using the invariance of the Haar measure and polar coordinates:

$$\frac{(\mathbb{E}_{F,Y}|P_F Y|^p)^{1/p}}{(\mathbb{E}_{F,Y}|P_F Y|^2)^{1/2}} = \frac{(\mathbb{E}_U h_{k,p}(U))^{1/p}}{(\mathbb{E}_U h_{k,2}(U))^{1/2}}, \quad (1.12)$$

where  $U$  is uniformly distributed over  $SO(n)$ ,  $E_0 \in G_{n,k}$ ,  $\theta_0 \in S(E_0)$ ,  $g$  denotes the density of  $Y$  in  $\mathbb{R}^n$ , and  $h_{k,q} : SO(n) \rightarrow \mathbb{R}_+$  is defined as:

$$h_{k,p}(u) := \text{Vol}(S^{k-1}) \int_0^\infty t^{p+k-1} \pi_{u(E_0)} g(tu(\theta_0)) dt. \quad (1.13)$$

To control the ratio in (1.12), a good bound on the log-Lipschitz constant  $L_{k,q}$  of  $h_{k,q}$  is required.

Our main technical result in this work is the following improvement over the log-Lipschitz bounds of Klartag from [19]:

**Theorem 1.5.** *Under the same assumptions as in Theorem 1.1,  $L_{k,p} \leq C \max(k,p)^{1/\alpha+1/2}$ .*

Contrary to Klartag's analytical approach for controlling the log-Lipschitz constant, ours is completely based on geometric convexity arguments, employing the convex bodies  $K_{k+q}$  introduced by K. Ball in [3], and a variation on the  $L_q$ -centroid bodies, which were introduced by E. Lutwak and G. Zhang in [22].

Fleury proceeds by employing three additional ingredients:

- As shown by Borell [9], for any log-concave function  $w$  on  $\mathbb{R}_+$ :

$$q \mapsto \log \frac{\int_0^\infty t^{q-1} w(t) dt}{\Gamma(q)} \text{ is concave on } \mathbb{R}_+ . \quad (1.14)$$

Consequently,  $q \mapsto \log(h_{k,q}(u)/\Gamma(k+q))$  is concave for any fixed  $u \in SO(n)$ . This ingredient was also used in [13].

- As follows e.g. from the work of Bakry and Émery [2] (see also [21]), for any Lipschitz function  $f : SO(n) \rightarrow \mathbb{R}_+$ , the following log-Sobolev inequality is satisfied (see Sections 2 and 3 for definitions):

$$\mathbb{E}nt_U(f) \leq \frac{c}{n} \mathbb{E}_U(|\nabla f|^2/f) . \quad (1.15)$$

- The latter log-Sobolev inequality implies via the Herbst argument, that for any log-Lipschitz function  $f : SO(n) \rightarrow \mathbb{R}_+$  with log-Lipschitz constant bounded above by  $L$ , the following reverse Hölder inequality holds (see [12, (15)]):

$$(\mathbb{E}_U f^q)^{\frac{1}{q}} \leq \exp\left(C \frac{L^2}{n}(q-r)\right) (\mathbb{E}_U f^r)^{\frac{1}{r}} \quad \forall q > r > 0 . \quad (1.16)$$

We proceed by using these ingredients as our predecessors, but our proof corrects the slight inefficiency of Fleury's approach in the resulting large-deviation estimate (witnessed by the comparison to Klartag's estimate earlier). The improvement here comes from the fact that we take the derivative in  $p$  of (1.11), and optimize on the dimension  $k$  for each  $p$  separately, as opposed to optimizing on a single  $k$  directly in (1.11). However, this by itself would not yield the improvement in the thin-shell estimate - the latter is due to our improved log-Lipschitz estimate in Theorem 1.5. Only by combining this improved log-Lipschitz estimate with our variation on Fleury's method, are we able to recover the sharp large-deviation estimates of Paouris (1.4).

The rest of this work is organized as follows. In Section 2 we prove a more general version of Theorem 1.5. In Section 3 we provide a complete proof of Theorem 1.2. In Section 4, we derive for completeness Theorem 1.1 from Theorem 1.2. In Section 5, we provide some concluding remarks. In the Appendix, we provide a proof of Proposition 2.6 and other lemmas, whose purpose is to handle the case when  $X$  is not centrally-symmetric (non-even density).

## 2 An improved log-Lipschitz estimate

Let  $M_{k,l}(\mathbb{R})$  denote the set of  $k$  by  $l$  matrices over  $\mathbb{R}$ . We equip

$$SO(n) = \{U \in M_{n,n}(\mathbb{R}); U^t U = Id, \det(U) = 1\}$$

with its standard (left and right) invariant Riemannian metric  $g$ , which we specify for concreteness on  $T_{Id}SO(n)$ , the tangent space at the identity element  $Id \in SO(n)$ .

Fixing an orthonormal basis of  $\mathbb{R}^n$  and taking the derivative of the relation  $U^t U = Id$ , we see that this tangent space may be identified with all anti-symmetric matrices  $\{B \in M_{n,n}(\mathbb{R}); B^t + B = 0\}$ . Given  $B \in T_{Id}SO(n)$ , we set  $|B|^2 := g_{Id}(B, B) = \frac{1}{2} \|B\|_{HS}^2$ , where recall the Hilbert-Schmidt norm of  $A \in M_{k,l}(\mathbb{R})$  is given by  $\|A\|_{HS}^2 := \text{tr}(A^t A) = \sum_{1 \leq i \leq k, 1 \leq j \leq l} A_{i,j}^2$ . The factor of  $\frac{1}{2}$  above is simply a convenience to ensure that a full  $2\pi$  degree rotation in any two-plane leaving the orthogonal complement in place, has geodesic length  $2\pi$ , and to prevent further appearances of factors like  $\sqrt{2}$  later on. Up to this factor, this metric coincides with the one induced from the natural embedding  $SO(n) \subset \mathbb{R}^{n^2}$ .

## 2.1 Main Result

Throughout this section, let  $Y$  denote an isotropic random vector in  $\mathbb{R}^n$  with log-concave density  $g$ . Given an integer  $k$  between 1 and  $n$ , a linear subspace  $E_0 \in G_{n,k}$  and  $\theta_0 \in S(E_0)$ , we recall the definition of the function  $h_{k,p} : SO(n) \rightarrow \mathbb{R}_+$ :

$$h_{k,p}(U) := \text{Vol}(S^{k-1}) \int_0^\infty t^{p+k-1} \pi_{U(E_0)} g(tU(\theta_0)) dt \quad , \quad U \in SO(n) . \quad (2.1)$$

Note that  $\pi_E g$  is log-concave for any  $E \in G_{n,k}$  by the Prékopa–Leindler theorem (e.g. [14]).

When  $Y = (X + G_n)/\sqrt{2}$ , where (as throughout this work)  $X$  denotes an isotropic random vector in  $\mathbb{R}^n$  with log-concave density, an upper bound on the log-Lipschitz constant (i.e. the Lipschitz constant of the logarithm) of:

$$U \mapsto \pi_{U(E_0)} g(tU(\theta_0))$$

was obtained by Klartag [19, Lemma 3.1], playing a crucial role in his polynomial estimates on the thin-shell of an isotropic log-concave measure. When  $t \leq C\sqrt{k}$ , Klartag's estimate is of the order of  $k^2$ . In [12], Fleury defined a truncated version of (2.1), where the integral ranges up to  $C\sqrt{k}$ . Klartag's estimate obviously implies the same bound on the log-Lipschitz constant of this truncated version of  $h_{k,p}$ .

Our main technical result in this work is the following improved estimate on the log-Lipschitz constant of  $h_{k,p}$ , which is completely based on geometric convexity arguments. Note that we do not need any truncation, nor do we need to assume that  $Y$  has been convolved with a Gaussian to obtain a meaningful estimate. However, the improvement over Klartag's  $k^2$  bound appears after this convolution.

**Theorem 2.1.** *The log-Lipschitz constant  $L_{k,p}$  of  $U \mapsto h_{k,p}(U)$  is bounded above by  $C \max(k, p) \text{dist}(Z_{\max(k,p)}^+(g), B_2^n)$ .*

Here  $Z_q^+(w) \subset \mathbb{R}^n$  ( $q \geq 1$ ) denotes the *one-sided*  $L_q$ -centroid body of the density  $w$  (which may not have total mass one), defined via its support functional:

$$h_{Z_q^+(w)}(y) = \left( 2 \int_{\mathbb{R}^n} \langle x, y \rangle_+^q w(x) dx \right)^{1/q} ,$$

(here as usual  $a_+ := \max(a, 0)$ ). When  $w$  is even, this coincides with the more standard definition of the  $L_q$ -centroid body, introduced by E. Lutwak and G. Zhang in [22] (under a different normalization):

$$h_{Z_q(w)}(y) = \left( \int_{\mathbb{R}^n} |\langle x, y \rangle|^q w(x) dx \right)^{1/q} .$$

Clearly:

$$Z_q^+(w) \subset 2^{1/q} Z_q(w) .$$

In any case, when  $w$  is the characteristic function of a set  $K$ , we denote  $Z_q^+(K) := Z_q^+(1_K)$ , and similarly for  $Z_q(K)$ . Lastly, the *geometric distance*  $\text{dist}(K, L)$  between two subsets  $K, L \subset \mathbb{R}^n$  is defined as:

$$\text{dist}(K, L) := \inf \{ C_2/C_1; C_1 L \subset K \subset C_2 L, C_1, C_2 > 0 \} .$$

A very useful result for handling the non-even case is due to Grünbaum [15]:

**Lemma 2.2.** *Let  $X_1$  denote a random variable on  $\mathbb{R}$  with log-concave density and barycenter at the origin. Then  $\frac{1}{e} \leq \mathbb{P}(X_1 \geq 0) \leq 1 - \frac{1}{e}$ .*

Note that by definition  $Y$  (and its density  $g$ ) is  $\psi_\alpha$  ( $\alpha \geq 1$ ) iff  $Z_q(g) \subset Cq^{1/\alpha} Z_2(g)$  for some fixed universal constant  $C > 1$  and all  $q \geq 2$ . Also recall that by Borell's Lemma [10], a log-concave probability density  $g$  is always  $\psi_1$ , and that moreover:

$$1 \leq q_1 \leq q_2 \quad \Rightarrow \quad Z_{q_1}(g) \subset Z_{q_2}(g) \subset C \frac{q_2}{q_1} Z_{q_1}(g) . \quad (2.2)$$

If in addition the barycenter of  $g$  is at the origin, then repeating the argument leading to (2.2) and using Lemma 2.2, one verifies:

$$1 \leq q_1 \leq q_2 \quad \Rightarrow \quad \left( \frac{2}{e} \right)^{\frac{1}{q_1} - \frac{1}{q_2}} Z_{q_1}^+(g) \subset Z_{q_2}^+(g) \subset C \left( \frac{2e-2}{e} \right)^{\frac{1}{q_1} - \frac{1}{q_2}} \frac{q_2}{q_1} Z_{q_1}^+(g) . \quad (2.3)$$

Note that  $Z_2(g) = B_2^n$  by definition of isotropicity, and one may similarly show (see Lemma A.5) that  $cB_2^n \subset Z_2^+(g) \subset \sqrt{2}B_2^n$ . It follows immediately from (2.3) that  $\text{dist}(Z_k^+(g), B_2^n) \leq Ck$ , and we see that Theorem 2.1 recovers Klartag's  $k^2$  order of magnitude when  $p \leq k$  (which is the case of interest in the subsequent analysis).

The improvement over Klartag's bound comes from the following elementary:

**Lemma 2.3.** *Let  $X_0$  denote a random-vector in  $\mathbb{R}^n$  with log-concave density and barycenter at the origin. Set  $Y_0 = (X_0 + G_n)/\sqrt{2}$  and denote by  $g_0$  its density. Then:*

1.  $Z_q^+(g_0) \supset c\sqrt{q}B_2^n$  for all  $q \geq 2$ .
2. If  $X_0$  is  $\psi_\alpha$  ( $\alpha \in [1, 2]$ ), then so is  $Y_0$ .





Denote by  $T_i$  ( $i = 1, 2, 3$ ) the subspace of  $T_{Id}SO(n)$  having the form (2.4) with  $V_j = 0$  for  $j \neq i$ . Given  $B \in T_i$ , we call the geodesic  $\mathbb{R} \ni s \mapsto U_s := \exp_{Id}(sB) \in SO(n)$  a *Type- $i$  movement*. Clearly  $\|V_i\|_{HS} = \sup_{0 \neq B \in T_i} \langle \nabla_{Id} \log h_{k,p}, B \rangle / |B|$ , so our goal now is to obtain a uniform upper bound on the derivative of  $\log h_{k,p}$  induced by a Type- $i$  movement.

To this end, we recall the following crucial fact, due to K. Ball [3, Theorem 5] in the even case, and verified to still hold in the general one by Klartag [17, Theorem 2.2]:

**Theorem.** *Let  $w$  denote a log-concave function on  $\mathbb{R}^m$  with  $0 < \int w < \infty$  and  $w(0) > 0$ . Given  $q \geq 1$ , set:*

$$\|x\| = \|x\|_{K_q(w)} := \left( q \int_0^\infty t^{q-1} w(tx) dt \right)^{-\frac{1}{q}}, \quad x \in \mathbb{R}^m.$$

Then for all  $x, y \in \mathbb{R}^m$ ,  $0 \leq \|x\| < \infty$ ,  $\|x\| = 0$  iff  $x = 0$ ,  $\|\lambda x\| = \lambda \|x\|$  for all  $\lambda \geq 0$ , and  $\|x + y\| \leq \|x\| + \|y\|$ .

We will thus say that  $\|\cdot\|_{K_q(w)}$  defines a norm, even though it may fail to be even, and denote by  $K_q(w) := \{x \in \mathbb{R}^m; \|x\|_{K_q(w)} \leq 1\}$  its associated convex compact unit-ball. Note that the constant  $q$  in front of the integral above is simply a convenient normalization for later use. We also set  $\|x\|_{\hat{K}_q(w)} := \max(\|x\|_{K_q(w)}, \|-x\|_{K_q(w)})$ , having unit-ball  $\hat{K}_q(w) = K_q(w) \cap -K_q(w)$ . Note that the triangle inequality implies that:

$$\left| \|x\|_{K_q(w)} - \|y\|_{K_q(w)} \right| \leq \|x - y\|_{\hat{K}_q(w)}. \quad (2.5)$$

Finally, note that since  $B_2^m$  is centrally-symmetric, then  $C_1 B_2^m \subset K \cap -K \subset K \subset C_2 B_2^m$  iff  $C_1 B_2^m \subset K \subset C_2 B_2^m$ , and hence:

$$\frac{\|x\|_{\hat{K}_q(w)}}{\|y\|_{K_q(w)}} \leq \text{dist}(K_q(w), B_2^m) \frac{|x|}{|y|}. \quad (2.6)$$

### 2.2.1 Type-1 movement

Let  $B \in T_1$  with  $|B| = 1$  generate a Type-1 movement  $\{U_s\}$ , and denote  $\xi_0 = \frac{d}{ds} U_s(\theta_0)|_{s=0} \in T_{\theta_0}S(\mathbb{R}^n)$ . Using henceforth the natural embedding  $T_{\theta}S(\mathbb{R}^n) \subset T_{\theta}\mathbb{R}^n \subset \mathbb{R}^n$ , we see that  $\xi_0$  lies in the orthogonal complement of  $\theta_0$  in  $E_0$ , and  $|B| = 1$  ensures that  $|\xi_0| = 1$ . So  $U_s$  is a rotation in the  $\{\theta_0, \xi_0\}$  plane, and  $U_s(E_0) = E_0$ . Recalling the definition of  $h_{k,p}$ , we see that:

$$|\langle \nabla_{Id} \log h_{k,p}, B \rangle| = \left| \frac{d}{ds} \log h_{k,p}(U_s) \Big|_{s=0} \right| = (k+p) \left| \frac{d}{ds} \log \|U_s(\theta_0)\|_{K_{k+p}(\pi_{E_0}g)} \Big|_{s=0} \right|.$$

Estimating the derivative of the norm using the triangle-inequality (2.5) and (2.6), we immediately obtain:

$$|\langle \nabla_{Id} \log h_{k,p}, B \rangle| \leq (k+p) \frac{\|\xi_0\|_{\hat{K}_{k+p}(\pi_{E_0}g)}}{\|\theta_0\|_{K_{k+p}(\pi_{E_0}g)}} \leq (k+p) \text{dist}(K_{k+p}(\pi_{E_0}g), B_2(E_0)).$$

### 2.2.2 Type-2 movement

Let  $B \in T_2$  with  $|B| = 1$  generate a Type-2 movement  $\{U_s\}$ , and denote  $\theta_s := U_s(\theta_0)$  and  $\xi_s := \frac{d}{ds}\theta_s \in T_{\theta_s}S(\mathbb{R}^n)$ . Now  $\xi_0 \in S(E_0^\perp)$ , and the rotation  $U_s$  in the  $\{\theta_s, \xi_s\} = \{\theta_0, \xi_0\}$  plane rotates  $E_0$  into  $E_s := U_s(E_0)$ . Denote  $H := \text{span}(E_s, \xi_s) \in G_{n,k+1}$ , and observe that:

$$h_{k,p}(U_s) = \text{Vol}(S^{k-1}) \int_0^\infty \int_{-\infty}^\infty t^{p+k-1} \pi_{Hg}(t\theta_s + r\xi_s) dr dt .$$

Performing the change of variables  $r = vt$ , which is valid except at the negligible point  $t = 0$ , we obtain:

$$h_{k,p}(U_s) = \text{Vol}(S^{k-1}) \int_0^\infty \int_{-\infty}^\infty t^{p+k} \pi_{Hg}(t(\theta_s + v\xi_s)) dv dt = c_{p,k} \int_{-\infty}^\infty \|\theta_s + v\xi_s\|_{K_{k+p+1}(\pi_{Hg})}^{-(k+p+1)} dv ,$$

where  $c_{p,k} = \text{Vol}(S^{k-1})/(k+p+1)$  is totally immaterial. Using that  $\frac{d}{ds}\xi_s = -\theta_s$  and the triangle inequality for  $\|\cdot\|_{K_{k+p+1}(\pi_{Hg})}$ , we obtain:

$$\begin{aligned} |\langle \nabla_{Id} \log h_{k,p}, B \rangle| &= \left| \frac{d}{ds} \log h_{k,p}(U_s) \Big|_{s=0} \right| \leq (k+p+1) \sup_{v \in \mathbb{R}} \frac{\|\xi_0 - v\theta_0\|_{\hat{K}_{k+p+1}(\pi_{Hg})}}{\|\theta_0 + v\xi_0\|_{K_{k+p+1}(\pi_{Hg})}} \\ &\leq (k+p+1) \text{dist}(K_{k+p+1}(\pi_{Hg}), B_2(H)) \sup_{v \in \mathbb{R}} \frac{|\xi_0 - v\theta_0|}{|\theta_0 + v\xi_0|} \\ &= (k+p+1) \text{dist}(K_{k+p+1}(\pi_{Hg}), B_2(H)) , \end{aligned}$$

where we have used the fact that  $\theta_0$  and  $\xi_0$  are orthogonal unit vectors in the last equality.

### 2.2.3 Type-3 movement

Finally, we analyze the most important movement type, which is responsible for a subspace of movements of dimension  $(k-1)(n-k)$  (out of the  $\dim G_{n,k} + \dim S^{k-1} = k(n-k) + (k-1)$  dimensional subspace of non-trivial movements).

Let  $0 \neq B \in T_3$  generate a Type-3 movement  $\{U_s\}$ , and note that  $U_s(\theta_0) = \theta_0$ . Set  $e_s^j := U_s(e^j)$ ,  $j = 2, \dots, k$ , and note that  $f^j := \frac{d}{ds}e_s^j|_{s=0} \in E_0^\perp$ . Denote  $F_0 := \text{span}\{f^2, \dots, f^k\}$ , and note that by slightly perturbing  $B$  if necessary, we may assume that  $F_0$  is  $k-1$  dimensional. Finally, set  $H = E_0 \oplus F_0 \in G_{n,2k-1}$ , and notice that  $H$  is invariant under  $U_s$  (since  $U_s$  is an isometry acting as the identity on the orthogonal complement). Consequently,  $H = E_s \oplus F_s$ , where  $F_s := U_s(F_0)$ , and so:

$$h_{k,p}(U_s) = \text{Vol}(S^{k-1}) \int_0^\infty \int_{F_s} t^{p+k-1} \pi_{Hg}(t\theta_0 + y) dy dt .$$

Using the change of variables  $y = zt$ , we obtain (with  $c_{p,k} = \text{Vol}(S^{k-1})/(2k-1+p)$ ):

$$h_{k,p}(U_s) = \text{Vol}(S^{k-1}) \int_0^\infty \int_{F_s} t^{p+2k-2} \pi_{Hg}(t(\theta_0 + z)) dz dt = c_{p,k} \int_{F_s} \|\theta_0 + z\|_{K_{2k-1+p}(\pi_{Hg})}^{-(2k-1+p)} dz ,$$

which we rewrite, since  $U_s$  is orthogonal, as:

$$h_{k,p}(U_s) = c_{p,k} \int_{F_0} \|\theta_0 + U_s(z)\|_{K_{2k-1+p}(\pi_{Hg})}^{-(2k-1+p)} dz .$$

As usual, the triangle inequality for  $\|\cdot\|_{K_{2k-1+p}(\pi_{Hg})}$  implies that:

$$|\langle \nabla_{Id} \log h_{k,p}, B \rangle| = \left| \frac{d}{ds} \log h_{k,p}(U_s) \Big|_{s=0} \right| \leq (2k-1+p) \sup_{z \in F_0} \frac{\|Bz\|_{\hat{K}_{2k-1+p}(\pi_{Hg})}}{\|\theta_0 + z\|_{K_{2k-1+p}(\pi_{Hg})}} ,$$

and so:

$$\begin{aligned} \frac{|\langle \nabla_{Id} \log h_{k,p}, B \rangle|}{(2k-1+p) \text{dist}(K_{2k-1+p}(\pi_{Hg}), B_2(H))} &\leq \sup_{z \in F_0} \frac{|Bz|}{|\theta_0 + z|} \leq \|B\|_{op} \sup_{z \in F_0} \frac{|z|}{\sqrt{1+|z|^2}} \\ &\leq \frac{\|B\|_{HS}}{\sqrt{2}} = |B| , \end{aligned}$$

where we have used that  $\theta_0$  is perpendicular to  $F_0$ , and that  $\|B\|_{op} \leq \|B\|_{HS}/\sqrt{2}$  for any anti-symmetric matrix  $B$  (here  $\|B\|_{op}$  denotes its operator norm), as may be easily verified by using the Cauchy–Schwarz inequality.

### 2.3 Distance of $K_{m+p}$ to Euclidean ball

To conclude the proof of Theorem 2.1, it remains to control the geometric distance of  $K_{m+p}(\pi_{Hg})$  to a Euclidean ball, for  $H \in G_{n,m}$  with  $m$  of the order of  $k$ . To this end, we compare  $K_{m+p}(\pi_{Hg})$  to  $Z_p(\pi_{Hg}) = P_H Z_p(g)$ . Our motivation comes from the work of Paouris [29], who noted that:

$$Z_p(\pi_{Hg}) = Z_p(K_{m+p}(\pi_{Hg})) ,$$

and using the trivial  $Z_p(K) \subset \text{conv}(K \cup -K)$  for any set  $K$  of volume 1, obtained an upper bound on  $\text{Vol}(Z_p(\pi_{Hg}))$  by bounding above  $\text{Vol}(K_{m+p}(\pi_{Hg}))$ . In this work, on the other hand, we take the converse path, passing from  $K_{m+p}$  bodies to  $Z_p$  ones, and consequently need to introduce the  $Z_p^+$  bodies to handle non-even densities. Moreover, we require bounds on  $Z_p^+(K)$  both from above and from below, which turn out to be more laborious in the non-even case (when  $K$  is not centrally-symmetric).

Since the distance to the Euclidean ball cannot increase under orthogonal projections, and since  $c_1 Z_k^+(g) \subset c_2 Z_m^+(g) \subset c_3 Z_{2k-1}^+(g) \subset c_4 Z_k^+(g)$  when  $k \leq m \leq 2k-1$  by (2.3), it remains to establish the following:

**Theorem 2.5.** *Let  $w$  denote a log-concave function on  $\mathbb{R}^m$  with  $0 < \int w < \infty$  and barycenter at the origin. Then for any  $p \geq 1$ :*

$$\text{dist}(K_{m+p}(w), B_2^m) \leq C \text{dist}(Z_{\max(p,m)}^+(w), B_2^m) .$$

For the proof, we recall several useful properties of the bodies  $K_q(w)$  and  $Z_q^+(K)$ . First, it is known (see [4, 3, 26] for the even case and [17, Lemmas 2.5,2.6] or [30, Lemma 3.2 and (3.12)] for the general one) that under the assumptions of Theorem 2.5:

$$1 \leq q_1 \leq q_2 \quad \Rightarrow \quad e^{-m(\frac{1}{q_1} - \frac{1}{q_2})} \frac{K_{q_1}(w)}{w(0)^{1/q_1}} \subset \frac{K_{q_2}(w)}{w(0)^{1/q_2}} \subset \frac{\Gamma(q_2 + 1)^{1/q_2}}{\Gamma(q_1 + 1)^{1/q_1}} \frac{K_{q_1}(w)}{w(0)^{1/q_1}}. \quad (2.7)$$

Second, integration in polar coordinates (cf. [29]) directly shows that:

$$Z_p^+(K_{m+p}(w)) = Z_p^+(w). \quad (2.8)$$

Lastly, we require the following proposition, which is well-known in the even-case (e.g. [28, Lemma 4.1]), but requires more work in the general one (note for instance that the barycenter of  $K_{m+p}(w)$  below need not be at the origin); its proof is postponed to the Appendix.

**Proposition 2.6.**

$$C_1 Z_p^+(K_{m+p}(w)) \subset \text{Vol}(K_{m+p}(w))^{1/p} K_{m+p}(w) \subset C_2 Z_p^+(K_{m+p}(w)) \left( \frac{\Gamma(m+p+1)}{\Gamma(m)\Gamma(p+1)} \right)^{1/p} \quad (2.9)$$

*Proof of Theorem 2.5.* Note that by Stirling's formula, (2.9) with (2.8) implies that:

$$\text{dist}(K_{m+p}(w), B_2^m) \leq C \frac{p+m}{p} \text{dist}(Z_p^+(w), B_2^m),$$

and so when  $p \geq m$  the asserted claim follows. Otherwise, using (2.7), Stirling's formula, (2.9) and (2.8), we see that if  $q \geq p$  then:

$$\text{dist}(K_{m+p}(w), B_2^m) \leq C_1 \frac{m+q}{m+p} \text{dist}(K_{m+q}(w), B_2^m) \leq C_2 \frac{m+q}{m+p} \frac{m+q}{q} \text{dist}(Z_q^+(w), B_2^m).$$

Setting  $q = m$ , the case  $p < m$  is also resolved.  $\square$

The proof of Theorem 2.1 is now complete.

### 3 Moment Estimates

In this section we provide a complete proof of Theorem 1.2.

#### 3.1 Reductions

Given  $X$  as in Theorem 1.2, set  $Y := (X + G_n)/\sqrt{2}$ , where  $G_n$  is an independent standard Gaussian random vector in  $\mathbb{R}^n$ . Note that  $Y$  is centrally-symmetric and isotropic, and by the Prékopa–Leindler Theorem, has log-concave density.

We repeat the argument of Fleury for reducing the moment estimation problem from  $X$  to  $Y$  and for passing from integration on  $\mathbb{R}^n$  to  $SO(n)$ . By the symmetry and independence of  $G_n$ , convexity of  $t \mapsto t^p$  and the Cauchy–Schwarz inequality, we have:

$$\begin{aligned} \mathbb{E}|Y|^{2p} &= E \left( \frac{|X + G_n|^2}{2} \right)^p = \frac{1}{2} \mathbb{E} \left( \left( \frac{|X + G_n|^2}{2} \right)^p + \left( \frac{|X - G_n|^2}{2} \right)^p \right) \\ &\geq \mathbb{E} \left( \frac{|X|^2 + |G_n|^2}{2} \right)^p \geq \mathbb{E}|X|^p |G_n|^p = \mathbb{E}|X|^p \mathbb{E}|G_n|^p \geq \mathbb{E}|X|^p (\mathbb{E}|G_n|^2)^{p/2} = n^{p/2} \mathbb{E}|X|^p . \end{aligned}$$

Since  $\mathbb{E}|X|^2 = \mathbb{E}|Y|^2 = n$ , we deduce:

$$\frac{(\mathbb{E}|X|^p)^{1/p}}{(\mathbb{E}|X|^2)^{1/2}} \leq \left( \frac{(\mathbb{E}|Y|^{2p})^{1/2p}}{(\mathbb{E}|Y|^2)^{1/2}} \right)^2 , \quad (3.1)$$

and it remains to obtain (1.9) with  $X$  replaced by  $Y$ , with an obvious modification of the constants.

Next, since  $|x|^p = a_{n,k,p} \mathbb{E}_F |P_F x|^p$ , where  $F$  is uniformly distributed on  $G_{n,k}$  (according to its Haar probability measure), with  $k$  to be determined later on, we have:

$$\frac{\mathbb{E}|Y|^p}{\mathbb{E}|G_n|^p} = \frac{\mathbb{E} \mathbb{E}_F |P_F Y|^p}{\mathbb{E} \mathbb{E}_F |P_F G_n|^p} = \frac{\mathbb{E} \mathbb{E}_F |P_F Y|^p}{\mathbb{E}|G_k|^p} ,$$

where  $G_i$  denotes a standard Gaussian random vector on  $\mathbb{R}^i$ . A direct calculation shows that:

$$\mathbb{E}|G_i|^p = 2^{p/2-1} \frac{\Gamma((p+i)/2)}{\Gamma(i/2)} ,$$

and hence:

$$\mathbb{E}|Y|^p = \frac{\Gamma((p+n)/2) \Gamma(k/2)}{\Gamma(n/2) \Gamma((p+k)/2)} \mathbb{E} \mathbb{E}_F |P_F Y|^p .$$

Passing to polar coordinates on  $F \in G_{n,k}$  and using the invariance of the Haar measures on  $G_{n,k}$ ,  $S(F)$  and  $SO(n)$  under the action of  $SO(n)$ , we verify that:

$$\mathbb{E} \mathbb{E}_F |P_F Y|^p = \mathbb{E}_U h_{k,p}(U) ,$$

where  $U$  is uniformly distributed on  $SO(n)$ .

### 3.2 Controlling the derivative

We now deviate from Fleury's argument and proceed to estimate:

$$\frac{d}{dp} \log((\mathbb{E}|Y|^p)^{\frac{1}{p}}) = \frac{d}{dp} \log((\mathbb{E}_U h_{k,p}(U))^{\frac{1}{p}}) + \frac{d}{dp} \left( \frac{1}{p} \log \frac{\Gamma((p+n)/2) \Gamma(k/2)}{\Gamma(n/2) \Gamma((p+k)/2)} \right) . \quad (3.2)$$

Given  $u \in SO(n)$ , we introduce the (non-probability) measure  $\mu_u$  on  $\mathbb{R}_+$  having density  $\text{Vol}(S^{k-1}) t^{k-1} \pi_{u(F_0)} g(tu(\theta_0))$ , where  $g$  is the density of  $Y$  on  $\mathbb{R}^n$ . We define the (probability) measure  $\mu_{k,p} := \mathbb{E}_U \mu_U$  on  $\mathbb{R}_+$ , and write:

$$h_{k,p}(u) = \mathbb{E}_{\mu_u}(t^p) , \quad \mathbb{E}_U h_{k,p}(U) = \mathbb{E}_U \mathbb{E}_{\mu_U}(t^p) = \mathbb{E}_{\mu_{k,p}}(t^p) .$$

Here and in the sequel we use the following convention: given a measure space  $(\Omega, \mu)$ , which does not necessarily have total mass 1, and a measurable  $f : \Omega \rightarrow \mathbb{R}_+$ , we set:

$$\mathbb{E}_\mu f = \mathbb{E}_\mu(f) = \int f d\mu \quad , \quad \mathbb{Ent}_\mu(f) = \mathbb{E}_\mu(f \log f) - \mathbb{E}_\mu(f) \log(\mathbb{E}_\mu(f)) \quad .$$

A useful fact, easily verified by direct calculation, is that:

$$\frac{d}{dp} \log((\mathbb{E}_\mu f^p)^{\frac{1}{p}}) = \frac{1}{p^2} \frac{\mathbb{Ent}_\mu(f^p)}{\mathbb{E}_\mu(f^p)} \quad .$$

We proceed with estimating (3.2). As explained:

$$\frac{d}{dp} \log((\mathbb{E}_U h_{k,p}(U))^{\frac{1}{p}}) = \frac{1}{p^2} \frac{\mathbb{Ent}_{\mu_{k,p}}(t^p)}{\mathbb{E}_{\mu_{k,p}}(t^p)} = \frac{1}{p^2} \frac{\mathbb{Ent}_{\mu_{k,p}}(t^p)}{\mathbb{E}_U h_{k,p}(U)} \quad . \quad (3.3)$$

Our main idea here is to decompose the numerator as follows:

$$\mathbb{Ent}_{\mu_{k,p}}(t^p) = \mathbb{E}_U \mathbb{Ent}_{\mu_U}(t^p) + \mathbb{Ent}_U \mathbb{E}_{\mu_U}(t^p) = \mathbb{E}_U \mathbb{Ent}_{\mu_U}(t^p) + \mathbb{Ent}_U h_{k,p}(U) \quad . \quad (3.4)$$

The contribution of the second term in (3.4) is controlled using the log-Sobolev inequality (1.15):

$$\frac{1}{p^2} \frac{\mathbb{Ent}_U h_{k,p}(U)}{\mathbb{E}_U h_{k,p}(U)} \leq \frac{c}{p^2 n} \frac{\mathbb{E}_U(|\nabla \log h_{k,p}|^2(U) h_{k,p}(U))}{\mathbb{E}_U h_{k,p}(U)} \leq \frac{c L_{k,p}^2}{p^2 n} \quad , \quad (3.5)$$

where recall  $L_{k,p}$  denotes the log-Lipschitz constant of  $u \mapsto h_{k,p}(u)$ . To control the contribution of the first term in (3.4), we first write given  $u \in SO(n)$ :

$$\frac{1}{p^2} \frac{\mathbb{Ent}_{\mu_u}(t^p)}{\mathbb{E}_{\mu_u}(t^p)} = \frac{d}{dp} \log((\mathbb{E}_{\mu_u} t^p)^{\frac{1}{p}}) = \frac{d}{dp} \frac{1}{p} \left( \log \frac{h_{k,p}(u)}{\Gamma(k+p)} - \log \frac{h_{k,0}(u)}{\Gamma(k)} + \log \frac{\Gamma(k+p)}{\Gamma(k)} + \log h_{k,0}(u) \right) \quad .$$

By Borell's concavity result (1.14), we realize that:

$$\frac{d}{dp} \frac{1}{p} \left( \log \frac{h_{k,p}(u)}{\Gamma(k+p)} - \log \frac{h_{k,0}(u)}{\Gamma(k)} \right) \leq 0 \quad ,$$

and hence:

$$\frac{1}{p^2} \frac{\mathbb{Ent}_{\mu_u}(t^p)}{\mathbb{E}_{\mu_u}(t^p)} \leq \frac{d}{dp} \left( \frac{1}{p} \log \frac{\Gamma(k+p)}{\Gamma(k)} \right) - \frac{1}{p^2} \log h_{k,0}(u) \quad .$$

Plugging this estimate back into (3.3) and (3.4), we obtain:

$$\frac{1}{p^2} \frac{\mathbb{E}_U \mathbb{Ent}_{\mu_U}(t^p)}{\mathbb{E}_U \mathbb{E}_{\mu_U}(t^p)} \leq \frac{d}{dp} \left( \frac{1}{p} \log \frac{\Gamma(k+p)}{\Gamma(k)} \right) + \frac{1}{p^2} \frac{\mathbb{E}_U \log(1/h_{k,0}(U)) h_{k,p}(U)}{\mathbb{E}_U h_{k,p}(U)} \quad . \quad (3.6)$$

By using the Jensen and Cauchy–Schwarz inequalities, we bound the second term by:

$$\frac{\mathbb{E}_U \log(1/h_{k,0}(U)) h_{k,p}(U)}{\mathbb{E}_U h_{k,p}(U)} \leq \log \left( \frac{\mathbb{E}_U \frac{h_{k,p}(U)}{h_{k,0}(U)}}{\mathbb{E}_U h_{k,p}(U)} \right) \leq \log \left( \frac{(\mathbb{E}_U h_{k,p}(U))^2}{\mathbb{E}_U h_{k,p}(U) (\mathbb{E}_U h_{k,0}(U))^{-2}} \right)^{1/2} \quad .$$

We now use the reverse Hölder inequality (1.16) for comparing the various moments above. Denoting  $\|f\|_q := (\mathbb{E}_U |f(U)|^q)^{1/q}$ , we have:

$$\|h_{k,p}\|_2 \leq \exp\left(\frac{CL_{k,p}^2}{n}\right) \|h_{k,p}\|_1 ,$$

$$\|h_{k,0}^{-1}\|_2 \leq \exp\left(\frac{2CL_{k,0}^2}{n}\right) \|h_{k,0}^{-1}\|_0 = \exp\left(\frac{2CL_{k,0}^2}{n}\right) \frac{1}{\|h_{k,0}\|_0} \leq \exp\left(\frac{3CL_{k,0}^2}{n}\right) \frac{1}{\|h_{k,0}\|_1} ,$$

where  $\|f\|_0$  is as usual interpreted as  $\exp(\mathbb{E}_U \log |f(U)|)$ . Since  $\|h_{k,0}\|_1 = \mathbb{E}_U h_{k,0}(U) = \mathbb{E}_{\mu_{k,p}}(1) = 1$ , we conclude that:

$$\frac{\mathbb{E}_U \log(1/h_{k,0}(U)h_{k,p}(U))}{\mathbb{E}_U h_{k,p}(U)} \leq \frac{C}{p^2 n} (L_{k,p}^2 + 3L_{k,0}^2) . \quad (3.7)$$

Now, plugging all the estimates (3.5), (3.6), (3.7) into (3.3) using the decomposition (3.4), and plugging the result into (3.2), we obtain:

$$\frac{d}{dp} \log((\mathbb{E}|Y|^p)^{\frac{1}{p}}) \leq \frac{c}{p^2 n} (2L_{k,p}^2 + 3L_{k,0}^2) + \frac{d}{dp} \left( \frac{1}{p} \log \frac{\Gamma(k+p)}{\Gamma(k)} \right) + \frac{d}{dp} \left( \frac{1}{p} \log \frac{\Gamma((p+n)/2)\Gamma(k/2)}{\Gamma(n/2)\Gamma((p+k)/2)} \right) .$$

### 3.3 Optimizing on the dimension

As observed by Fleury, using that the function  $\frac{d}{dp} \log \Gamma(p)$  is concave, the contribution of the last term above is easily verified to be non-positive and moreover insignificant relative to the second term, so we just bound it from above by 0. For the second term, we estimate using Jensen's inequality, for any  $q > 0$ :

$$\begin{aligned} \frac{d}{dp} \left( \frac{1}{p} \log \frac{\Gamma(k+p)}{\Gamma(k)} \right) &= \frac{1}{pq} \frac{\int_0^\infty \log(t^q) t^{p+k-1} \exp(-t) dt}{\Gamma(p+k)} - \frac{1}{p^2} \log \frac{\Gamma(k+p)}{\Gamma(k)} \\ &\leq \frac{1}{pq} \log \frac{\Gamma(k+p+q)}{\Gamma(k+p)} - \frac{1}{p^2} \log \frac{\Gamma(k+p)}{\Gamma(k)} = \frac{1}{p} \log \left( \frac{\Gamma(k+p+q)^{1/q}}{\Gamma(k+p)^{1/q}} \frac{\Gamma(k)^{1/p}}{\Gamma(k+p)^{1/p}} \right) . \end{aligned}$$

Applying Stirling's formula and setting  $q = (p+k-1)\frac{p}{k-1}$ , one verifies that:

$$\frac{d}{dp} \left( \frac{1}{p} \log \frac{\Gamma(k+p)}{\Gamma(k)} \right) \leq \frac{C}{k} .$$

Plugging our estimates for  $L_{k,q}$  obtained in Corollary 2.4, we conclude that if  $X$  is  $\psi_\alpha$  ( $\alpha \in [1, 2]$ ), then:

$$\frac{d}{dp} \log((\mathbb{E}|Y|^p)^{\frac{1}{p}}) \leq C \left( \frac{k^{1+2/\alpha}}{p^2 n} + \frac{1}{k} \right) \quad \forall k \in [p, n] .$$



Optimizing on  $k$  in the above range, we set:

$$k = \lceil p^{1/\beta} n^{1/(2\beta)} \rceil, \quad \beta := 1 + \frac{1}{\alpha},$$

which is guaranteed to satisfy  $k \in [p, n]$  whenever  $2 \leq p \leq n^{\alpha/2}$ , and obtain for such  $p$ :

$$\frac{d}{dp} \log((E|Y|^p)^{\frac{1}{p}}) \leq \frac{C_2}{p^{1/\beta} n^{1/(2\beta)}}.$$

Integrating over  $p$ , we obtain for  $p$  in that range:

$$(\mathbb{E}|Y|^p)^{\frac{1}{p}} \leq \exp\left(C_3 \frac{p^{\frac{1}{\alpha+1}}}{n^{\frac{\alpha}{2(\alpha+1)}}}\right) (\mathbb{E}|Y|^2)^{\frac{1}{2}},$$

and together with the reduction (3.1) from  $X$  to  $Y$ , the conclusion of Theorem 1.2 follows.

## 4 Deviation Estimates

Obtaining the deviation estimates of Theorem 1.1 from the moment estimates of Theorem 1.2 is completely standard, exactly as in [12]. For completeness, we provide a brief description.

*Proof of Theorem 1.1.* Set:

$$\varepsilon_{n,\alpha} := \min\left(1, \frac{2^{\frac{\alpha+2}{\alpha+1}} C}{n^{\frac{\alpha}{2(\alpha+1)}}}\right),$$

and note that there exists a constant  $t_0 \in (0, 1]$ , so that:

$$\forall t \in [\varepsilon_{n,\alpha}, t_0] \quad \exists p \in [2, cn^{\alpha/2}] \quad \text{such that} \quad t = 2C \frac{p^{\frac{1}{\alpha+1}}}{n^{\frac{\alpha}{2(\alpha+1)}}}. \quad (4.1)$$

Here  $c, C > 0$  are the two constants appearing in Theorem 1.2, which guarantee that for  $p$  in the above range:

$$(\mathbb{E}|X|^p)^{1/p} \leq \left(1 + \frac{t}{2}\right) \sqrt{n}.$$

Since  $\frac{1+t}{1+t/2} \geq 1 + t/3$  for  $t \in [0, 1]$ , we obtain by the Markov–Chebyshev inequality:

$$\mathbb{P}(|X| \geq (1+t)\sqrt{n}) \leq \mathbb{P}(|X| \geq (1+t/3)(\mathbb{E}|X|^p)^{1/p}) \leq (1+t/3)^{-p} \leq \exp(-pt/4).$$

Expressing  $p$  as a function of  $t$ , for  $t$  in the range specified in (4.1), and plugging this above, we obtain:

$$\mathbb{P}(|X| \geq (1+t)\sqrt{n}) \leq \exp(-c_1 n^{\alpha/2} t^{2+\alpha}) \quad \forall t \in [\varepsilon_{n,\alpha}, t_0].$$

To extend this estimate to the entire interval  $[0, t_0]$ , note that:

$$\mathbb{P}(|X| \geq (1+t)\sqrt{n}) \leq (1+t)^{-2} \leq \exp(-t/2) \quad \forall t \in [0, \varepsilon_{n,\alpha}] ,$$

and so adjusting the constants appearing above:

$$\mathbb{P}(|X| \geq (1+t)\sqrt{n}) \leq \exp(-c_2 n^{\alpha/2} t^{2+\alpha}) \quad \forall t \in [0, t_0] .$$

Finally, a standard application of Borell's lemma [10] (e.g. as in [29]), ensures that:

$$\mathbb{P}(|X| \geq (1+t)\sqrt{n}) \leq \exp(-c_3 n^{\alpha/2} t) \quad \forall t \geq t_0 ,$$

concluding the proof of the positive deviation estimate (1.6).

For the proof of the negative deviation estimate (1.7), observe that there exists a constant  $c_4 > 0$ , so that setting  $p_0 := c_4 n^{\frac{\alpha}{2(\alpha+2)}}$ , Theorem 1.2 implies that:

$$\mathbb{E}|X|^{2p_0} \leq \left(1 + C \frac{p_0^{\frac{1}{\alpha+1}}}{n^{\frac{\alpha}{2(\alpha+1)}}}\right)^{2p_0} (\mathbb{E}|X|^2)^{p_0} \leq \frac{17}{16} (\mathbb{E}|X|^2)^{p_0} \leq \frac{17}{16} (\mathbb{E}|X|^{p_0})^2 .$$

Consequently  $\text{Var}|X|^{p_0} \leq \frac{1}{16} (\mathbb{E}|X|^{p_0})^2$ , and Chebyshev's inequality implies:

$$\begin{aligned} \frac{1}{4} &\geq \mathbb{P}\left(\left||X|^{p_0} - \mathbb{E}|X|^{p_0}\right| \geq \frac{1}{2} \mathbb{E}|X|^{p_0}\right) \\ &\geq \mathbb{P}\left(|X| \geq \frac{1}{2^{1/p_0}} (\mathbb{E}|X|^{p_0})^{1/p_0}\right) \geq \mathbb{P}\left(|X| \leq \left(1 - \frac{c_5}{p_0}\right) \sqrt{n}\right) . \end{aligned} \quad (4.2)$$

On the other hand, the positive deviation estimate (1.6) implies that:

$$\mathbb{P}\left(|X| \leq \left(1 + \frac{c_6}{p_0}\right) \sqrt{n}\right) \geq \frac{3}{4} . \quad (4.3)$$

Setting  $t_1 := \frac{\max(c_5, c_6)}{p_0}$  and given  $t \in (t_1, 1)$ , we set  $\lambda := \frac{2t_1}{t+t_1} \in (0, 1)$  so that:

$$(1 - t_1) = \lambda(1 - t) + (1 - \lambda)(1 + t_1) ,$$

and by the log-concavity of the function  $\mathbb{R}_+ \ni s \mapsto \mathbb{P}(|X| \leq s)$  (a consequence of Prékopa–Leindler), it follows that:

$$\mathbb{P}(|X| \leq (1 - t_1)\sqrt{n}) \geq \mathbb{P}(|X| \leq (1 - t)\sqrt{n})^\lambda \mathbb{P}(|X| \leq (1 + t_1)\sqrt{n})^{1-\lambda} .$$

Using the estimates (4.2) and (4.3), we deduce:

$$\mathbb{P}(|X| \leq (1 - t)\sqrt{n}) \leq \left(\frac{1}{4}\right)^{1/\lambda} \left(\frac{4}{3}\right)^{1/\lambda-1} \leq \frac{1}{4} \frac{1}{3^{1/\lambda}} \leq \exp(-c_7 p_0 t) \quad \forall t \in (t_1, 1) .$$

The negative deviation estimate (1.7) immediately follows.

Lastly, we observe that:

$$\sqrt{\text{Var}|X|} \leq C \frac{n^{1/2}}{p_0} = C_2 n^{\frac{1}{2+\alpha}} ,$$

e.g. by integrating by parts and using the positive and negative deviation estimates (see e.g. [11, Lemma 6]).  $\square$

## 5 Concluding Remarks

**Remark 5.1.** Examining the proof, it is easy to verify that if the log-Lipschitz constant  $L_{k,p}$  of  $h_{k,p} : SO(n) \rightarrow \mathbb{R}_+$  satisfies:

$$p \leq k \quad \Rightarrow \quad L_{k,p} \leq Cp^\beta k^\gamma, \quad \beta, \gamma \in \mathbb{R},$$

then the sharp large-deviation estimate  $\mathbb{P}(|X| \geq C\sqrt{n}) \leq \exp(-\sqrt{n})$  is recovered if and only if  $\beta + \gamma = 3/2$ . Of course, since  $p \leq k$ , it is better to have larger  $\beta$ , and this affects the resulting thin-shell estimate. Our estimates yield  $\beta = 0$  and  $\gamma = 3/2$ .

**Remark 5.2.** Using a theorem of Bobkov [6], we improve the best-known bound on the Cheeger constant  $D_{Che}(\mu)$  of a general log-concave isotropic measure  $\mu$  in  $\mathbb{R}^n$  to  $D_{Che}(\mu) \geq cn^{-\frac{5}{12}}$ , bringing us a little bit closer to the full KLS conjecture  $D_{Che}(\mu) \geq c > 0$  (we refer to [6] for missing definitions and background). Note that the estimate improves all the way to  $D_{Che}(\mu) \geq cn^{-\frac{3}{8}}$  when  $\mu$  is  $\psi_2$ .

## Appendix

In the Appendix, we prove several properties of the bodies  $Z_p^+(K)$  which are needed for the results of Section 2.

Our main goal is to establish Proposition 2.6. Given  $\theta \in S^{m-1}$ , we denote  $H_\theta^+ := \{x \in \mathbb{R}^m; \langle x, \theta \rangle \geq 0\}$  and set  $H_\theta^- := H_{-\theta}^+$ . For the proof, we require several lemmas.

**Lemma A.1.** *Let  $K$  denote a convex body in  $\mathbb{R}^m$ , and given  $\theta \in S^{m-1}$ , denote  $f_\theta = \pi_\theta \mathbf{1}_K$ . Then:*

$$\left( \frac{f_\theta(0)}{\|f_\theta\|_\infty} \right)^{1/p} \left( \frac{\Gamma(m)\Gamma(p+1)}{\Gamma(m+p+1)} \right)^{1/p} h_K(\theta) \leq \frac{h_{Z_p^+(K)}(\theta)}{(2\text{Vol}(K \cap H_\theta^+))^{1/p}} \leq h_K(\theta).$$

*Proof.* The right inequality is straightforward from the definitions. The left inequality is derived by following the proof of [28, Lemma 4.1], which uses the fact that the  $1/(m-1)$  power of any one-dimensional marginal of  $K$  is a concave function.  $\square$

To control the left-most term in Lemma A.1, we have:

**Lemma A.2.** *Let  $\mu = f(x)dx$  denote a log-concave probability measure on  $\mathbb{R}$ . Then for any  $\varepsilon > 0$ :*

$$\varepsilon \leq \int_0^\infty f(x)dx \leq 1 - \varepsilon \quad \Rightarrow \quad f(0) \geq \varepsilon \|f\|_\infty.$$

*Proof.* Consider the function  $\mathcal{I} : [0, 1] \rightarrow \mathbb{R}_+$  given by  $\mathcal{I}(v) = \min(f \circ F^{-1}(v), f \circ F^{-1}(1-v))$ , where  $F(x) = \int_{-\infty}^x f(t)dt$ . By a result of Bobkov [5],  $\mathcal{I}$  is the isoperimetric profile of the measure-metric space  $(\mathbb{R}, |\cdot|, \mu)$  (see [5] for definitions), and furthermore,  $\mathcal{I}$  is a

concave and symmetric function on  $[0, 1]$ . If  $v = F(0)$  is such that  $v \in [\varepsilon, 1 - \varepsilon]$ , it follows by the concavity of  $\mathcal{I}$  that:

$$f(0) \geq \mathcal{I}(v) \geq 2 \min(v, 1 - v) \mathcal{I}(1/2) \geq 2\varepsilon f(m) , \quad (\text{A.1})$$

where  $m = F^{-1}(1/2)$  is the median of  $\mu$ . But by [24, Lemma 2.7],  $f(m) \geq \|f\|_\infty / 2$ , which together with (A.1) concludes the assertion.  $\square$

This reduces our task to showing:

**Lemma A.3.** *If  $w$  is a log-concave function on  $\mathbb{R}^m$  with barycenter at the origin, then:*

$$\forall \theta \in S^{m-1} \quad \frac{1}{C} \leq \left( \frac{\text{Vol}(K_{m+p}(w) \cap H_\theta^+)}{\text{Vol}(K_{m+p}(w) \cap H_\theta^-)} \right)^{1/p} \leq C .$$

*Proof.* Note that we may normalize and rescale so that  $w(0) = 1$  and  $\int_{\mathbb{R}^m} w(x) dx = 1$ . Using polar-coordinates, we have for any convex (in fact, star-shaped) body  $K$  containing the origin:

$$\text{Vol}(K \cap H_\theta^+) = \frac{1}{m} \int_{S^{m-1} \cap H_\theta^+} \|\xi\|_K^{-m} d\xi . \quad (\text{A.2})$$

Using (2.7), we see that:

$$\forall \xi \in S^{m-1} \quad e^{-\frac{mp}{m+p}} \|\xi\|_{K_m(w)}^{-m} \leq \|\xi\|_{K_{m+p}(w)}^{-m} \leq \frac{\Gamma(m+p+1)^{\frac{m}{m+p}}}{\Gamma(m+1)} \|\xi\|_{K_m(w)}^{-m} .$$

Plugging this into (A.2) and using Stirling's formula, we verify that:

$$\forall \theta \in S^{m-1} \quad e^{-p} \leq \frac{\text{Vol}(K_{m+p}(w) \cap H_\theta^+)}{\text{Vol}(K_m(w) \cap H_\theta^+)} \leq C^p . \quad (\text{A.3})$$

Using (A.2), the definition of  $K_m(w)$  and polar-coordinates again, we see that  $\text{Vol}(K_m(w) \cap H_\theta^+) = \int_{H_\theta^+} w(x) dx = \mathbb{P}(W_1 \geq 0)$ , where  $W_1$  is the random variable on  $\text{span}(\theta)$  having density  $\pi_\theta w$ . Since this density is log-concave by the Prékopa–Leindler theorem, and since the barycenter of  $W_1$  is at the origin, Lemma 2.2 implies that:

$$\frac{1}{e-1} \leq \frac{\text{Vol}(K_m(w) \cap H_\theta^+)}{\text{Vol}(K_m(w) \cap H_\theta^-)} \leq e-1 .$$

Together with (A.3), this concludes the proof.  $\square$

**Corollary A.4.** *With the same assumptions as in Lemma A.3:*

$$\forall \theta \in S^{m-1} \quad \frac{1}{C'} \leq \left( \frac{\text{Vol}(K_{m+p}(w) \cap H_\theta^+)}{\text{Vol}(K_{m+p}(w))} \right)^{1/p} \leq C' .$$

*Proof of Proposition 2.6.* Applying Lemma A.1 with  $K = K_{m+p}(w)$  and using Corollary A.4, we obtain for all  $\theta \in S^{m-1}$ :

$$c \left( \frac{f_\theta(0)}{\|f_\theta\|_\infty} \right)^{1/p} \left( \frac{\Gamma(m)\Gamma(p+1)}{\Gamma(m+p+1)} \right)^{1/p} \leq \text{Vol}(K_{m+p}(w))^{-1/p} \frac{h_{Z_p^+(K_{m+p}(w))}(\theta)}{h_{K_{m+p}(w)}(\theta)} \leq C .$$

Lemma A.2 together with Lemma A.3 imply that:

$$\forall \theta \in S^{m-1} \quad \left( \frac{f_\theta(0)}{\|f_\theta\|_\infty} \right)^{1/p} \geq c' > 0 ,$$

and hence:

$$c' \left( \frac{\Gamma(m)\Gamma(p+1)}{\Gamma(m+p+1)} \right)^{1/p} K_{m+p}(w) \subset \text{Vol}(K_{m+p}(w))^{-1/p} Z_p^+(K_{m+p}(w)) \subset CK_{m+p}(w) .$$

Rearranging terms, the assertion of Proposition 2.6 follows.  $\square$

Finally, we prove:

**Lemma A.5.** *If  $g : \mathbb{R}^m \rightarrow \mathbb{R}_+$  is log-concave and isotropic then  $Z_2^+(g) \supset cB_2^m$ .*

*Proof.* Given  $\theta \in S^{n-1}$ , denote  $g_\theta := \pi_\theta g$ ; as usual, it is an isotropic log-concave probability density on  $\mathbb{R}$ . Comparing moments using (2.7) with  $m = 1$ ,  $q_1 = 1$  and  $q_2 = 3$ , we obtain:

$$3 \int_0^\infty t^2 g_\theta(t) dt \geq \frac{\left( \int_0^\infty g_\theta(t) dt \right)^3}{e^2 g_\theta(0)^2} .$$

Applying now the reverse comparison in both directions  $\theta$  and  $-\theta$  and summing, we obtain:

$$3 = 3 \int_{-\infty}^\infty t^2 g_\theta(t) dt \leq \frac{\Gamma(4)}{g_\theta(0)^2} \left( \left( \int_0^\infty g_\theta(t) dt \right)^3 + \left( \int_{-\infty}^0 g_\theta(t) dt \right)^3 \right) .$$

Combining these two estimates and using Lemma 2.2 to control  $\int_0^\infty g_\theta(t) dt$ , the assertion follows with e.g.  $c = (3e^2(1 + (e-1)^3))^{-1/2}$ .  $\square$

## References

- [1] M. Anttila, K. Ball, and I. Perissinaki. The central limit problem for convex bodies. *Trans. Amer. Math. Soc.*, 355(12):4723–4735, 2003.
- [2] D. Bakry and M. Émery. Diffusions hypercontractives. In *Séminaire de probabilités, XIX, 1983/84*, volume 1123 of *Lecture Notes in Math.*, pages 177–206. Springer, Berlin, 1985.

- [3] K. Ball. Logarithmically concave functions and sections of convex sets in  $\mathbb{R}^n$ . *Studia Math.*, 88(1):69–84, 1988.
- [4] R. E. Barlow, A. W. Marshall, and F. Proschan. Properties of probability distributions with monotone hazard rate. *Ann. Math. Statist.*, 34:375–389, 1963.
- [5] S. Bobkov. Extremal properties of half-spaces for log-concave distributions. *Ann. Probab.*, 24(1):35–48, 1996.
- [6] S. Bobkov. On isoperimetric constants for log-concave probability distributions. In *Geometric aspects of functional analysis, Israel Seminar 2004-2005*, volume 1910 of *Lecture Notes in Math.*, pages 81–88. Springer, Berlin, 2007.
- [7] S. G. Bobkov and A. Koldobsky. On the central limit property of convex bodies. In *Geometric aspects of functional analysis*, volume 1807 of *Lecture Notes in Math.*, pages 44–52. Springer, Berlin, 2003.
- [8] S. G. Bobkov and F. L. Nazarov. On convex bodies and log-concave probability measures with unconditional basis. In *Geometric Aspects of Functional Analysis*, volume 1807 of *Lecture Notes in Mathematics*, pages 53–69. Springer, 2001-2002.
- [9] Ch. Borell. Complements of Lyapunov’s inequality. *Math. Ann.*, 205:323–331, 1973.
- [10] Ch. Borell. Convex measures on locally convex spaces. *Ark. Mat.*, 12:239–252, 1974.
- [11] B. Fleury. Between Paouris concentration inequality and variance conjecture. *Ann. Inst. Henri Poincaré Probab. Stat.*, 46(2):299–312, 2010.
- [12] B. Fleury. Concentration in a thin euclidean shell for log-concave measures. *J. Func. Anal.*, 259:832–841, 2010.
- [13] B. Fleury, O. Guédon, and G. Paouris. A stability result for mean width of  $l_p$ -centroid bodies. *Advances in Mathematics*, 214(2):865–877, 2007.
- [14] R. J. Gardner. The Brunn-Minkowski inequality. *Bull. Amer. Math. Soc. (N.S.)*, 39(3):355–405, 2002.
- [15] B. Grünbaum. Partitions of mass-distributions and of convex bodies by hyperplanes. *Pacific J. Math.*, 10:1257–1261, 1960.
- [16] R. Kannan, L. Lovász, and M. Simonovits. Isoperimetric problems for convex bodies and a localization lemma. *Discrete Comput. Geom.*, 13(3-4):541–559, 1995.
- [17] B. Klartag. On convex perturbations with a bounded isotropic constant. *Geom. and Funct. Anal.*, 16(6):1274–1290, 2006.
- [18] B. Klartag. A central limit theorem for convex sets. *Invent. Math.*, 168:91–131, 2007.

- [19] B. Klartag. Power-law estimates for the central limit theorem for convex sets. *J. Funct. Anal.*, 245:284–310, 2007.
- [20] B. Klartag. A Berry-Esseen type inequality for convex bodies with an unconditional basis. *Probab. Theory Related Fields*, 45(1):1–33, 2009.
- [21] M. Ledoux. *The concentration of measure phenomenon*, volume 89 of *Mathematical Surveys and Monographs*. American Mathematical Society, Providence, RI, 2001.
- [22] E. Lutwak and G. Zhang. Blaschke-Santaló inequalities. *J. Differential Geom.*, 47(1):1–16, 1997.
- [23] E. Milman. On gaussian marginals of uniformly convex bodies. *J. Theoret. Prob.*, 22(1):256–278, 2009.
- [24] E. Milman and S. Sodin. An isoperimetric inequality for uniformly log-concave measures and uniformly convex bodies. *J. Funct. Anal.*, 254(5):1235–1268, 2008.
- [25] V. D. Milman. A new proof of A. Dvoretzky’s theorem on cross-sections of convex bodies. *Funkcional. Anal. i Priložen.*, 5(4):28–37, 1971.
- [26] V. D. Milman and A. Pajor. Isotropic position and interia ellipsoids and zonoids of the unit ball of a normed  $n$ -dimensional space. In *Geometric Aspects of Functional Analysis*, volume 1376 of *Lecture Notes in Mathematics*, pages 64–104. Springer-Verlag, 1987-1988.
- [27] V. D. Milman and G. Schechtman. *Asymptotic theory of finite-dimensional normed spaces*, volume 1200 of *Lecture Notes in Mathematics*. Springer-Verlag, Berlin, 1986. With an appendix by M. Gromov.
- [28] G. Paouris.  $\psi_2$ -estimates for linear functionals on zonoids. In *Geometric Aspects of Functional Analysis*, volume 1807 of *Lecture Notes in Mathematics*, pages 211–222. Springer, 2001-2002.
- [29] G. Paouris. Concentration of mass on convex bodies. *Geom. Funct. Anal.*, 16(5):1021–1049, 2006.
- [30] G. Paouris. Small ball probability estimates for log-concave measures. To appear in *Trans. Amer. Math. Soc.*, 2010.