

MULTIPLICATION OF FREE RANDOM VARIABLES AND THE S -TRANSFORM: THE CASE OF VAN- ISHING MEAN

N. RAJ RAO¹

*Massachusetts Institute of Technology, Department of Mathematics, 77 Massachusetts Avenue,
Cambridge, MA 02139. USA.*

email: raj@mit.edu

ROLAND SPEICHER²

*Queen's University, Department of Mathematics and Statistics, Jeffery Hall, Kingston, ON
K7L 3N6. Canada.*

email: speicher@mast.queensu.ca

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Abstract

This note extends Voiculescu's S -transform based analytical machinery for free multiplicative convolution to the case where the mean of the probability measures vanishes. We show that with the right interpretation of the S -transform in the case of vanishing mean, the usual formula makes perfectly good sense.

1 Introduction

Multiplicative free convolution \boxtimes was introduced by Voiculescu [7], as an operation on probability measures to describe the multiplication of free random variables. Since various classes of random matrices become asymptotically free, the multiplicative free convolution is an important concept to deal with the asymptotic eigenvalue distribution of products of random matrices [8, 1]. A powerful analytical tool for an effective calculation of \boxtimes is Voiculescu's S -transform. In the usual presentations of the S -transform it looks as if its definition breaks down in the case where the considered random variable has mean zero. Since interesting cases, like the multiplicative free convolution of a semicircle with a free Poisson distribution (corresponding in the random matrix language to the product of a Gaussian random matrix with an independent Wishart matrix), fall into this class one gets the impression that the S -transform

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machinery has a crucial flaw here. However, as we want to point out, this is only ostensive, and with the right interpretation of the *S*-transform in the case of vanishing mean, the usual formula makes perfectly good sense. We will show this by combinatorial arguments in the setting where all moments exists. It stands to reason that one should be able to establish this using purely analytical tools, for more general distributions for which the combinatorial moment based arguments do not suffice.

2 Main result: one mean is zero

Let us first recall the definition of the *S*-transform. We start by working on an algebraic level, further below we will address the question of positivity (which is crucial for the definition of \boxtimes). So a random variable x is an element in some unital algebra \mathcal{A} , which is equipped with a linear functional φ such that $\varphi(1) = 1$. Our main interest is in the moments $\varphi(x^n)$ of the random variable x . Of course, of main interest for us are real-valued random variables $X : \Omega \rightarrow \mathbb{R}$ in the genuine sense, where φ is given by taking the expectation with respect to the underlying probability measure,

$$\varphi(X^n) = \int_{\Omega} X(\omega)^n dP(\omega).$$

In our algebraic frame we have to restrict to situations where moments determine the distribution uniquely; let us, however, remark that by analytic tools much of the theory of multiplicative free convolution can be extended to classes of probability measures with no assumption on the existence of moments.

Definition 2.1. Let x be a random variable with $\varphi(x) \neq 0$. Then its *S*-transform S_x is defined as follows. Let χ denote the inverse under composition of the series

$$\psi(z) := \sum_{n=1}^{\infty} \varphi(x^n)z^n,$$

then

$$S_x(z) := \chi(z) \cdot \frac{1+z}{z}.$$

The above definition has to be understood on the level of formal power series. Note that $\varphi(x) \neq 0$ ensures that the inverse of ψ exists as formal power series in z . Furthermore, ψ can be recovered from S_x and thus the *S*-transform contains all information about the moments of the considered random variable.

The relevance of the *S*-transform in free probability theory is due to the following theorem of Voiculescu [7].

Theorem 2.2. *If x and y are free random variables such that $\varphi(x) \neq 0$ and $\varphi(y) \neq 0$, then we have*

$$S_{xy}(z) = S_x(z) \cdot S_y(z).$$

Note that in the above setting $\varphi(xy) = \varphi(x)\varphi(y) \neq 0$, thus the *S*-transform of xy is also well-defined and the above theorem allows to get the moments of xy out of the moments of x and the moments of y .

Note that in this theorem we do not require x or y to be real or positive random variables. In this formulation, the theorem is true in full generality, since it is essentially a statement

about mixed moments of free random variables. However, if one wants to extract a probability distribution out of the S -transform, then one has to make special requirements concerning x and y . In particular, one would like to start with a situation where $\varphi(x^n)$ are the moments of a probability measure μ on \mathbb{R} (i.e., x should be a real-valued random variable, i.e., a selfadjoint operator) and where $\varphi(y^n)$ are the moments of a probability measure ν on \mathbb{R} (i.e., y should be real-valued random variable or selfadjoint operator). Since x and y do not commute, xy will not be a selfadjoint operator, thus it is not clear (and in general will not be the case) that $\varphi((xy)^n)$ are the moments of a probability measure. However, if x is a positive random variable, then \sqrt{x} makes sense and the operator $\sqrt{xy}\sqrt{x}$ is selfadjoint and has the same moments as xy . Thus in this case we know that the $\varphi((xy)^n)$ are the moments of a probability measure on \mathbb{R} ; this is called the *multiplicative free convolution of μ and ν* and denoted by $\mu \boxtimes \nu$. So the operation $\mu, \nu \mapsto \mu \boxtimes \nu$ is defined for probability measures μ, ν on \mathbb{R} if at least one of them is supported on the positive real line.

Often one restricts both factors to be supported in \mathbb{R}_+ . In this case the above theorem for the S -transform gives a solution to the problem of calculating $\mu \boxtimes \nu$. But there are also many interesting cases where only one factor is supported on \mathbb{R}_+ . In this case the description of $\mu \boxtimes \nu$ in terms of the S -transform seems to have the following flaw. In order to be able to invert the power series $\psi(z) = \sum_{n=1}^{\infty} \varphi(x^n)z^n$ one needs a non-vanishing linear term, which means that the mean $\varphi(x)$ of the variable should be non-zero. For measures supported on \mathbb{R}_+ this is (with the exception of the uninteresting case $x = 0$) satisfied, but for measures supported on \mathbb{R} the mean $\varphi(x)$ might be zero and in such a case it seems that the S -transform cannot be used to calculate $\mu \boxtimes \nu$. One should note that this is not an artificial situation, but covers such common cases like taking the multiplicative free convolution of a semicircle with a distribution supported on \mathbb{R}_+ . One could of course try to approximate mean zero situations with non mean zero ones; however, the main purpose of our note is to point out that actually the failure of the S -transform description in the mean zero case is just ostensive; with the right interpretation, the above theorem can also be used in that case to determine $\mu \boxtimes \nu$.

So let us consider the situation that $\varphi(x) = 0$ for some selfadjoint element. Let us first see whether we still can make sense out of the relation $S_{xy}(z) = S_x(z) \cdot S_y(z)$ and whether this still allows to recover all moments of xy uniquely out of the moments of x and of y . Again we exclude the uninteresting case that $x = 0$; hence we know that $\varphi(x^2) > 0$ (otherwise, by positivity, x would be equal to 0) and thus our series $\psi(z)$ starts with a multiple of z^2 . This means that although it cannot be inverted by a power series in z it can be inverted by a power series in \sqrt{z} . This inverse is not unique, but there are actually two choices (which correspond to choosing a branch of \sqrt{z}); however, this ambiguity does not affect the final result for the moments of xy , if y has non-zero mean. Let us be more specific about this in the following proposition.

Proposition 2.3. (1) *Let ψ be a formal power series of the form*

$$\psi(z) = \sum_{n=2}^{\infty} \alpha_n z^n = \alpha_2 z^2 + \alpha_3 z^3 + \dots$$

with $\alpha_2 \neq 0$. Then there exist exactly two power series in \sqrt{z} which satisfy

$$\psi(\chi(z)) = z.$$

If we denote these solutions by

$$\chi(z) = \sum_{k=1}^{\infty} \beta_k z^{k/2} = \beta_1 \sqrt{z} + \beta_2 z + \beta_3 \sqrt{z}^3 + \dots$$

and

$$\tilde{\chi}(z) = \sum_{k=1}^{\infty} \tilde{\beta}_k z^{k/2} = \tilde{\beta}_1 \sqrt{z} + \tilde{\beta}_2 z + \tilde{\beta}_3 \sqrt{z}^3 + \dots$$

then their coefficients are related by

$$\tilde{\beta}_k = (-1)^k \beta_k.$$

(2) Let ψ be as above and consider the two corresponding S-transforms

$$S(z) = \chi(z) \cdot \frac{1+z}{z}$$

and

$$\tilde{S}(z) = \tilde{\chi}(z) \cdot \frac{1+z}{z}$$

Then $S(z)$ and $\tilde{S}(z)$ are of the form

$$S(z) = \gamma_{-1} \frac{1}{\sqrt{z}} + \sum_{k=0}^{\infty} \gamma_k z^{k/2}$$

and

$$\tilde{S}(z) = \tilde{\gamma}_{-1} \frac{1}{\sqrt{z}} + \sum_{k=0}^{\infty} \tilde{\gamma}_k z^{k/2},$$

where

$$\tilde{\gamma}_k = (-1)^k \gamma_k.$$

(3) Let x be a random variable with vanishing mean and denote by S_x and \tilde{S}_x the corresponding two S-transforms, as constructed in (1) and (2). Let y be another random variable, with non-vanishing mean so that its S-transform is of the form

$$S_y(z) = \sum_{k=0}^{\infty} \delta_k z^k = \delta_0 + \delta_1 z + \delta_2 z^2 + \dots$$

Then the two series $S_{xy}(z) := S_x(z)S_y(z)$ and $\tilde{S}_{xy}(z) := \tilde{S}_x(z)S_y(z)$ are of the form

$$S_{xy}(z) = \varepsilon_{-1} \frac{1}{\sqrt{z}} + \sum_{k=0}^{\infty} \varepsilon_k z^{k/2}$$

and

$$\tilde{S}_{xy}(z) = \tilde{\varepsilon}_{-1} \frac{1}{\sqrt{z}} + \sum_{k=0}^{\infty} \tilde{\varepsilon}_k z^{k/2}$$

and related by

$$\tilde{\varepsilon}_k = (-1)^k \varepsilon_k.$$

If we put

$$\chi_{xy}(z) = S_{xy}(z) \cdot \frac{z}{1+z}, \quad \tilde{\chi}_{xy}(z) = \tilde{S}_{xy}(z) \cdot \frac{z}{1+z}$$

and denote by ψ_{xy} the unique solution of

$$\psi_{xy}(\chi_{xy}(z)) = z$$

and by $\tilde{\psi}_{xy}$ the unique solution of

$$\tilde{\psi}_{xy}(\tilde{\chi}_{xy}(z)) = z,$$

then we have

$$\psi_{xy}(z) = \tilde{\psi}_{xy}(z).$$

Proof. (1) The equation $\psi(\chi(z)) = z$ means

$$z = \sum_{n=2}^{\infty} \alpha_n \left(\sum_{k=1}^{\infty} \beta_k z^{k/2} \right)^n = \sum_{n=2}^{\infty} \sum_{\substack{k_1, \dots, k_n=1 \\ k_1 + \dots + k_n = n}} \alpha_n \beta_{k_1} \cdots \beta_{k_n} z^{(k_1 + \dots + k_n)/2}.$$

By equating the coefficients of powers of \sqrt{z} this is equivalent to the system of equations

$$1 = \alpha_2 \beta_1^2$$

and

$$0 = \sum_{n=2}^r \sum_{\substack{k_1, \dots, k_n=1 \\ k_1 + \dots + k_n = r}} \alpha_n \beta_{k_1} \cdots \beta_{k_n}$$

for all $r > 2$.

The first equation has the two solutions

$$\beta_1 = \frac{1}{\sqrt{\alpha_2}}$$

and

$$\tilde{\beta}_1 = -\frac{1}{\sqrt{\alpha_2}} = -\beta_1.$$

Writing the other equations in the form

$$0 = 2\alpha_2 \beta_{r-1} \beta_1 + \sum_{n=2}^r \sum_{\substack{k_1, \dots, k_n=1 \\ k_1 + \dots + k_n = r \\ \dots}} \alpha_n \beta_{k_1} \cdots \beta_{k_n}$$

(where the \dots indicate that we exclude in the second sum the cases $k_1 = 1, k_2 = r - 1$ and $k_1 = r - 1, k_2 = 1$) one sees that the values of β_n for $n > 1$ are recursively determined by β_1 and the α 's. By induction, it follows that $\tilde{\beta}_1 = -\beta_1$ results in $\tilde{\beta}_k = (-1)^k \beta_k$ for all k .

(2) This is clear, since we have $\gamma_k = \beta_{k+2} + \beta_k$ for all $k = -1, 0, 1, \dots$, where we set $\beta_{-1} := \beta_0 := 0$.

(3) This follows by reverting the arguments from the first and second part.

□

Definition 2.4. Let x be a random variable with $\varphi(x) = 0$ and $\varphi(x^2) \neq 0$. Then its two S -transforms S_x and \tilde{S}_x are defined as follows. Let χ and $\tilde{\chi}$ denote the two inverses under composition of the series

$$\psi(z) := \sum_{n=1}^{\infty} \varphi(x^n)z^n = \varphi(x^2)z^2 + \varphi(x^3)z^3 + \dots,$$

then

$$S_x(z) := \chi(z) \cdot \frac{1+z}{z} \quad \text{and} \quad \tilde{S}_x(z) := \tilde{\chi}(z) \cdot \frac{1+z}{z}.$$

Both S_x and \tilde{S}_x are formal series in \sqrt{z} of the form

$$\gamma_{-1} \frac{1}{\sqrt{z}} + \sum_{k=0}^{\infty} \gamma_k z^{k/2}$$

Now we stand ready to formulate our main theorem

Theorem 2.5. Let x and y be free random variables such that $\varphi(x) = 0$, $\varphi(x^2) \neq 0$ and $\varphi(y) \neq 0$. By S_x and \tilde{S}_x we denote the two S -transforms of x . Then

$$S_{xy}(z) = S_x(z) \cdot S_y(z) \quad \text{and} \quad \tilde{S}_{xy}(z) = \tilde{S}_x(z) \cdot S_y(z)$$

are the two S -transforms of xy .

Note that xy falls into the realm of Def. 2.4. We have

$$\varphi(xy) = \varphi(x)\varphi(y) = 0$$

and

$$\varphi((xy)^2) = \varphi(x^2)\varphi(y)^2 + \varphi(x)^2\varphi(y^2) - \varphi(x)^2\varphi(y)^2 = \varphi(x^2)\varphi(y)^2 \neq 0.$$

Let us point out again that the ambiguity of having two S -transforms for xy is not a problem, because both of them will lead to the same moments for xy , according to Prop. 2.3.

In order to prove our Theorem 2.5, one might first inspect the usual proofs for the S -transform. One notices that all of them rely on the fact that $\varphi(x)$ and $\varphi(y)$ both are not zero. None of the published proofs can be used directly for the case where one variable is centered. However, at least the idea of the proof in [3, 4] can be adapted to our situation. The following proof is a variant of the approach in [3, 4] (by avoiding an explicit use of the incomplete boxed convolution) and can of course also be used to give a straightforward proof in the usual case $\varphi(x) \neq 0$.

Proof. For a random variable x we denote by

$$\psi_x(z) := \sum_{n=1}^{\infty} \varphi(x^n)z^n$$

and

$$M_x(z) := 1 + \psi_x(z)$$

the corresponding moments series and by

$$C_x(z) := \sum_{n=1}^{\infty} \kappa_n^x z^n$$

the corresponding cumulants series ($\kappa_n^x = \kappa_n(x, \dots, x)$ is here the n -th free cumulant of x). Consider now x and y as in the theorem. Then, in addition to the moment and cumulant series of x , y , and xy , we consider also additional moments series of the form

$$M_1(x) := \sum_{n=0}^{\infty} \varphi(y(xy)^n) z^n$$

and

$$M_2(z) := \sum_{n=0}^{\infty} \varphi(x(yx)^n) z^n.$$

By the moment cumulant formula and by the fact that because of the freeness between x and y mixed cumulants in x and y vanish (see [4] for the combinatorial theory of freeness), it is quite straightforward to derive the following relations between these power series:

$$M_{xy}(z) = M_{yx}(z) = C_y[zM_2(z)] + 1 \quad (1)$$

$$M_1(z) = C_y[zM_2(z)] \cdot \frac{M_{xy}(z)}{zM_2(z)} \quad (2)$$

$$M_2(z) = C_x[zM_1(z)] \cdot \frac{M_{xy}(z)}{zM_1(z)} \quad (3)$$

Note that all these relations are valid (and make sense as formal power series) independent of whether $\varphi(x) = 0$ or not.

Let us note that we have the well-known relation $M(z) = 1 + C[zM(z)]$ between moment and cumulant series, which shows, by replacing z with $\chi(z)$, that $C[zS(z)] = z$. This relation, which was observed in [3], can be taken as an alternate definition of $S(z)$. In the case where $\varphi(y) \neq 0$ this is just a relation between formal power series in z . In the case $\varphi(x) = 0$, it has again to be read as formal power series in \sqrt{z} . (Note that for the case $\varphi(x) = 0$, $\varphi(x^2) \neq 0$ we also have $\kappa_1^x = \varphi(x) = 0$ and $\kappa_2^x = \varphi(x^2) \neq 0$, thus $C_x(z)$ starts also with a quadratic term in z .)

Since $\varphi(xy) = \varphi(x)\varphi(y) = 0$, the moment series ψ_{xy} has two inverses; let χ be one of them. Replacing z by $\chi(z)$ in (1) gives

$$1 + z = M_{xy}(\chi(z)) = C_y[\chi(z)M_2(\chi(z))] + 1,$$

yielding that

$$z = C_y[\chi(z)M_2(\chi(z))]. \quad (4)$$

Equation (2) gives

$$zM_1(z)M_2(z) = C_y[zM_2(z)] \cdot M_{xy}(z).$$

Replacing z by $\chi(z)$ gives

$$\chi(z) \cdot M_1(\chi(z)) \cdot M_2(\chi(z)) = C_y[\chi(z)M_2(\chi(z))] \cdot M_{xy}(\chi(z))$$

and hence

$$\chi(z) \cdot M_1(\chi(z)) \cdot M_2(\chi(z)) = z \cdot (z + 1) \quad (5)$$

Since $zS_y(z)$ is the unique inverse of C_y under composition we must have, by (4) that

$$zS_y(z) = \chi(z)M_2(\chi(z)).$$

In the same way as for (4) we get

$$z = C_x[\chi(z)M_1(\chi(z))].$$

Since C_x has exactly two power series in \sqrt{z} as inverses under composition, namely $zS_x(z)$ and $z\tilde{S}_x(z)$ we must have either

$$zS_x(z) = \chi(z)M_1(\chi(z)) \quad \text{or} \quad z\tilde{S}_x(z) = \chi(z)M_1(\chi(z)).$$

Name S_x and \tilde{S}_x in such a way that we have the first equation. Plugging this into (5) gives

$$zS_x(z) \cdot zS_y(z) = \chi(z)M_1(\chi(z)) \cdot \chi(z)M_2(\chi(z)) = \chi(z)z(z + 1)$$

and thus

$$S_x(z)S_y(z) = \chi(z)\frac{z + 1}{z} = S_{xy}(z).$$

If we start with the other inverse, $\tilde{\chi}$, of ψ_{xy} , then we would end with the other S -transform, \tilde{S}_{xy} . □

This proposition tells us that we can also use the formula $S_{xy}(z) = S_x(z)S_y(z)$ to calculate the moment series of xy , for x and y free, in the case where at most one of them has non-vanishing mean. Since this is the case in the situation where y has a distribution supported on \mathbb{R}_+ , we see that the S -transform is a useful tool for the calculation of $\mu \boxtimes \nu$, whenever μ is a probability measure on \mathbb{R} and ν is a probability measure on \mathbb{R}_+ , independent of whether μ has vanishing mean or not.

3 Case where both means are zero

One should also note that for the proof of our proposition it is important that at least one of the two involved measures has non-vanishing mean. If both have vanishing mean then the arguments break down. One should, however, note that this cannot be attributed to the fact that there does not exist a probability measure with the moments of xy . Actually, for x, y free with $\varphi(x) = 0 = \varphi(y)$ we have by the definition of freeness that $\varphi((xy)^n) = 0$ for all n , i.e., xy has the same moments as the delta distribution δ_0 . So one might say that the multiplicative free convolution of any two measures with mean zero is δ_0 . However, on a formal level the situation for the S -transform gets quite different here. As a concrete example, let us consider the multiplicative free convolution of the semicircle with some ν . The S -transform of the semicircle μ (of variance 1) is

$$S_\mu(z) = \frac{1}{\sqrt{z}}.$$

If we multiply this with an S -transform which is a power series in z then the result is again a series of the form

$$\frac{1}{\sqrt{z}} \cdot \text{power series in } \sqrt{z},$$

which can, as in our proposition, be uniquely resolved for the corresponding moment series. Let us take now, on the other hand, the multiplicative convolution of two semicirculars. Then we have

$$S_\mu(z)S_\mu(z) = \frac{1}{z}$$

which results in a ψ transform of the form

$$\psi(z) = \frac{1}{z} - 1.$$

This, however, is not a valid moment series. On the other hand, δ_0 would have the moment series $\psi(z) = 0$, for which no corresponding S -transform exists. So we see that in the case where we multiply two free operators with mean zero the S -transform machinery breaks totally down.

4 Some examples

Consider the free multiplicative convolution of the semi-circle distribution with the free Poisson distribution. The semi-circle distribution has S -transform

$$S_\mu(z) = \frac{1}{\sqrt{z}}$$

while the free Poisson distribution has S -transform

$$S_\gamma(z) = \frac{1}{z+1}.$$

The distribution obtained by their free multiplicative convolution has S -transform

$$S_{\mu \boxtimes \gamma}(z) = \frac{1}{\sqrt{z}(z+1)}.$$

The Cauchy transform of the probability measure $\mu \boxtimes \gamma$ satisfies the algebraic equation

$$g^4 z^2 - zg + 1 = 0,$$

from which we can obtain the density function, plotted in Figure 1(a).

Consider now the free multiplicative convolution of the free Poisson distribution with the shifted (by $-\alpha$) free Poisson distribution. The free Poisson distribution (shifted by $-\alpha$) has S -transform

$$S_\mu^\alpha(z) = \frac{-z - 1 + \alpha + \sqrt{z^2 + 2z + 2z\alpha + 1 - 2\alpha + \alpha^2}}{2z\alpha}$$

which, for $\alpha = 1$ yields

$$S_\mu(z) := S_\mu^1(z) = \frac{-z + \sqrt{z^2 + 4z}}{2z}$$

The distribution obtained by the free multiplicative convolution of the free Poisson distribution with the free Poisson distribution (shifted by -1) has S -transform

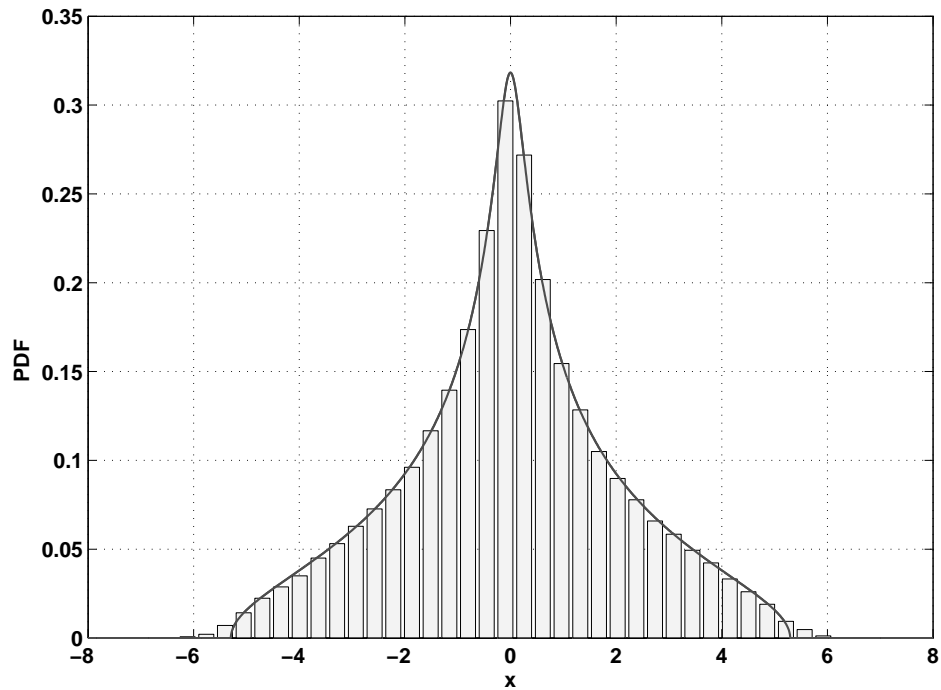
$$S_{\mu \boxtimes \gamma}(z) = \frac{-z + \sqrt{z^2 + 4z}}{2z(1+z)}.$$

The Cauchy transform of the probability measure $\mu \boxtimes \gamma$ satisfies the algebraic equation

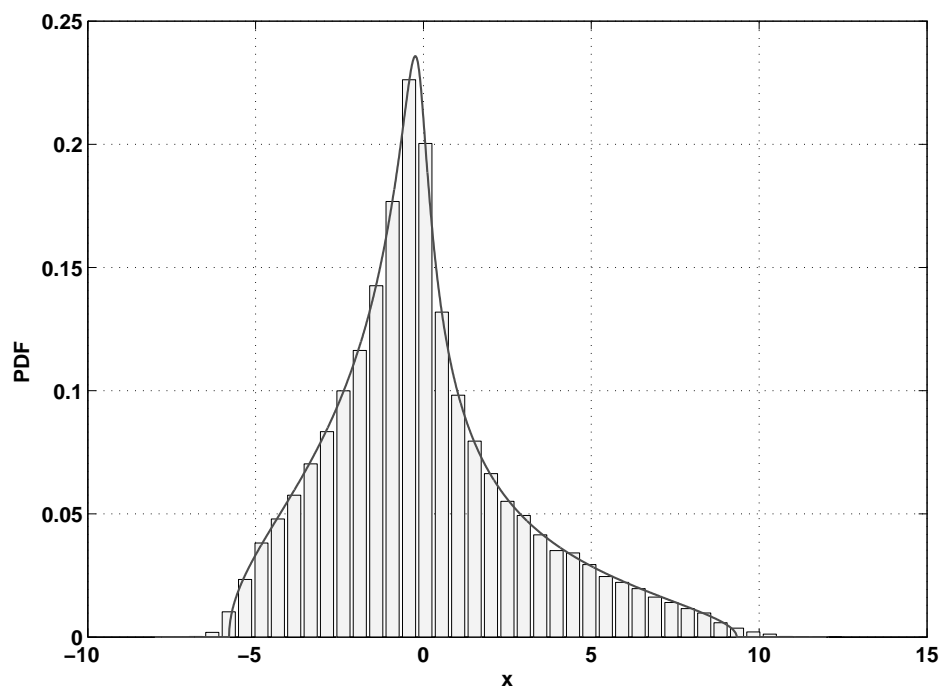
$$g^4 z^2 + z^2 g^3 - zg^2 - gz + 1 = 0$$

from which we can obtain the density function, plotted in Figure 1(b).

We note that the computations in this section were done using `RMTool` software [5] based on the algebraic equation based framework for computational free probability developed in [2, 6].



(a) The solid line is the density function of the probability measure obtained by the free multiplicative convolution of the semicircle with the free Poisson distribution. The histogram bars are the eigenvalues of the product of a Wigner matrix with a Wishart matrix and were generated using 50×50 sized random matrices over 4000 Monte-Carlo trials.



(b) The solid line is the density function of the probability measure obtained by the free multiplicative convolution of the free Poisson distribution with the shifted (by -1) free Poisson distribution. The histogram bars are the eigenvalues of the product of a Wishart matrix with an independent Wishart matrix (shifted by the negative identity) and were generated using 50×50 sized random matrices over 4000 Monte-Carlo trials.

Figure 1: Examples of free multiplicative convolution with vanishing mean.

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References

- [1] F. HIAI AND D. PETZ, *The semicircle law, free random variables and entropy*, vol. 77, American Mathematical Society, 2000. MR1746976
- [2] R. R. NADAKUDITI, *Applied Stochastic Eigen-Analysis*, PhD thesis, Massachusetts Institute of Technology, February 2007. Department of Electrical Engineering and Computer Science.
- [3] A. NICA AND R. SPEICHER, *A "Fourier transform" for multiplicative functions on non-crossing partitions*, *J. Algebraic Combin.*, 6 (1997), pp. 141–160. MR1436532
- [4] A. NICA AND R. SPEICHER, *Lectures on the Combinatorics of Free Probability*, London Mathematical Society Lecture Note Series, New York: Cambridge University Press, 2006.
- [5] N. R. RAO, *RMTool: A random matrix and free probability calculator in MATLAB*. <http://www.mit.edu/~raj/rmtool/>.
- [6] N. R. RAO AND A. EDELMAN, *The polynomial method for random matrices*. <http://arxiv.org/math.PR/0601389>, 2006.
- [7] D. VOICULESCU, *Multiplication of certain noncommuting random variables*, *J. Operator Theory*, 18 (1987), pp. 223–235. MR0915507
- [8] D. VOICULESCU, K. DYKEMA, AND A. NICA, *Free random variables*, vol. 1 of CRM Monograph Series, American Mathematical Society, Providence, RI, 1992. MR1217253